Baseline fibroblast growth factor 23 is associated with long-term mortality in ST-elevation myocardial infarction—results from the Augsburg myocardial infarction registry

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Background: The aim of this study was to investigate the association between inflammatory plasma protein concentrations and long-term mortality in patients with ST-elevation myocardial infarction (STEMI).

Methods: For 343 STEMI patients recorded between 2009 and 2013 by the population-based Myocardial Infarction Registry Augsburg, 92 inflammatory plasma proteins were measured at the index event using the OLINK inflammation panel. In multivariable-adjusted Cox regression models, the association between each plasma protein and all-cause long-term mortality was investigated. Median follow-up time was 7.6 (IQR: 2.4) years. For plasma protein that showed a strong association with long-term mortality, a 5-year survival ROC analysis was performed.

Results: One plasma protein, namely Fibroblast Growth Factor 23 (FGF-23), was particularly well associated with long-term mortality in the multivariable-adjusted Cox model with an FDR-adjusted p-value of <0.001 and a Hazard Ratio (HR) of 1.57 [95% CI: 1.29–1.91]. In the 5-years ROC analysis, an AUC of 0.6903 [95% CI: 0.594–0.781] was estimated for FGF-23. All other plasma protein didn’t show strong associations, each marker with FDR-adjusted p-values >0.05 in the multivariable-adjusted Cox models.

Conclusions: FGF-23 is independently associated with long-term mortality after STEMI and might play an important role in the response to myocardial injury. The results suggest FGF-23 to be a useful marker in the long-term treatment of STEMI patients and a potential target for drug development.

Keywords
FGF-23, inflammatory plasma protein, myocardial infarction, STEMI, long-term mortality
1. Background

Several prior studies investigated the association between a variety of inflammatory biomarkers and mortality after acute myocardial infarction (AMI) (1–6). Nevertheless, most studies concentrated on short-term morality or short-term major adverse cardiac events (MACE). There is a lack of evidence on whether inflammatory plasma proteins included in proteomics approaches, are associated with long-term outcomes in AMI patients. Thus, the aim of this study was to examine the association between a range of inflammation-related plasma proteins and long-term all-cause mortality in ST-elevation myocardial infarction (STEMI) patients and to identify the plasma proteins with the highest predictive ability for long-term mortality after the acute event.

2. Methods

2.1. Study population

This study is based on data from the population-based Augsburg Myocardial Infarction Registry. It was established in 1984 as a part of the MONICA-project (Monitoring Trends and Determinants in Cardiovascular disease) and operated as KORA Myocardial Infarction Registry (7) until 2019. The study area consists of the city of Augsburg, Germany, and the two adjacent counties comprising a total of approximately 680,000 inhabitants. Patients aged between 25 and 84 years being admitted for an AMI to one out of eight hospitals in the study area were consecutively registered. More detailed information on case identification, diagnostic classification of events and quality control of the data can be found in previous publications (7, 8). For the present study, only patients with AMI admitted to the university hospital of Augsburg between 2009 and 2014 were included, since blood from AMI patients was collected solely in this hospital. All study participants gave written informed consent. Methods of data collection have been approved by the ethics committee of the Bavarian Medical Association (Bayerische Landesaerztekammer) and the study was performed in accordance with the Declaration of Helsinki. The Augsburg Myocardial Infarction Registry is registered at “Deutsches Register Klinischer Studien” (DRKS) with the project number DRKS00029042.

2.1.1. Data collection

Trained study nurses interviewed the participants during hospital stay using a standardized questionnaire. In order to confirm the information provided by the patients and to collect additional information, the patients’ medical chart was reviewed. Demographic data, data on cardiovascular risk factors, medical history, comorbidities (including diabetes) and medication before and during hospital stay, as well as at discharge were collected from each patient. Furthermore, routine laboratory parameters including glucose measurement, ECG results and data on the in-hospital course were assessed.

Between 2009 and 2014 arterial EDTA blood samples were obtained immediately after hospital admission, right at the beginning of the cardiac catheterization. After centrifugation and aliquoting in the catheter laboratory, EDTA plasma samples were stored frozen at −80°C until laboratory analysis.

2.1.2. Clinical chemistry measurement

Measurements of the 92 proteins were performed in plasma samples using the Proseek® Multiplex Inflammation panel, developed by Olink Proteomics (Upplands Vägmans, Sweden) and based on the Proximity Extension Assay (PEA) technology. Detailed information on the process of measurement can be found in a previous publication (9) and directly at the website of Olink Proteomics (10).

All other blood parameters were measured in clinical laboratory at the university hospital of Augsburg during hospital stay of the patients as part of the regular diagnosis and routine treatment.

2.1.3. Outcome

The endpoint used in this study was long-term all-cause mortality. Mortality was ascertained by regularly checking the vital status of all registered persons of the MI registry with data from the population registries. Death certificates were obtained from local health departments. The present analysis included patients with ST elevation myocardial infarction (STEMI) only. All patients who died within the first 28 days after AMI were excluded in order to concentrate on long-term mortality exclusively. Overall, 343 STEMI patients survived the first 28 days after infarction and were considered for this analysis.

2.2. Statistical analysis

For the comparison of categorical variables, χ² tests were performed and the results were presented as absolute frequencies with percentages. For normally-distributed continuous variables, Student’s t-tests were used. For continuous variables that were not normally-distributed we used nonparametric tests. The results were presented as mean and standard deviation (SD) or median and inter-quartiles range (IQR).

2.2.1. Cox regression models

First, the obtained values for each plasma protein were standardized in order to ensure comparability between the different markers. To examine the associations between the 92 plasma proteins and long-term mortality, 92 Cox regression models were calculated, one for each protein. The initial models were adjusted for sex and age. To control the effect of multiple testing, the obtained p-values were false discovery rate (FDR)-adjusted. In a subsequent step, the same Cox regression models were calculated but potential confounder variables were added according to a literature review. These multivariable-adjusted Cox regression models were adjusted for sex, age, glomerular filtration...
rate (GFR; continuous variable), diabetes mellitus, hypertension, hyperlipidemia and percutaneous coronary intervention (PCI), yes/no).

The proportional hazards assumption was checked by plotting the Schoenfeld residuals against time and searching for any visible correlation and by testing for a significant correlation of the Schoenfeld residuals with time. The proportional hazards assumption was considered to be sufficiently fulfilled for all models.

As a sensitivity analysis, two more multivariable adjusted Cox regression models were calculated, each including the same covariables as the main model but adding either prehospital time or left ventricular ejection fraction (LVEF) (see Supplementary Figures S1, S2). Finally, in these models, it was tested for a significant interaction between the plasma proteins and prehospital or LVEF time, respectively, by including an interaction term into the Cox models.

2.2.2. ROC analysis, Youden-Index, and Kaplan-Meier curve

For plasma proteins that were strongly associated with long-term mortality ROC analyses was performed to quantify their predictive ability.

### 3. Results

For a total of 343 STEMI cases data of 92 inflammation-related plasma proteins were available for the statistical analysis. Table 1 displays the baseline characteristics of the total sample and stratified for patients who died within the observational period and those who didn’t.

As for most plasma proteins 3 values were missing, a total of 340 cases was included into the Cox regression models (there were no missing values for the covariables included into the main regression models). Results of the Cox regression models adjusted for sex and age are displayed in Figure 1. One protein, Fibroblast Growth Factor 23 (FGF-23), was strikingly associated with long-term mortality. Adjusting the Cox regression model for sex, age, GFR, diabetes mellitus, hypertension, hyperlipidemia and percutaneous coronary intervention (PCI), yes/no).

Best cut-off value for a dichotomization of the plasma protein markers were determined by calculating Youden-Indices. The sample then was divided in a "high plasma protein" group and a "low plasma protein" group according to the maximized Youden-Index. Finally, a Kaplan-Meier curve was plotted and stratified by the dichotomous variable high/low plasma protein group.

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**TABLE 1 Baseline characteristics for the total sample and stratified for event/no event (event = death 28 days after AMI and within observational period).**

<table>
<thead>
<tr>
<th></th>
<th>Total sample (n = 343)</th>
<th>No event (n = 257)</th>
<th>Event (n = 86)</th>
<th>p-value</th>
<th>n</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (mean, SD)</td>
<td>62.8 (11.8)</td>
<td>60.1 (11.1)</td>
<td>70.8 (10.1)</td>
<td>&lt;0.001</td>
<td>343</td>
</tr>
<tr>
<td>Male</td>
<td>252 (73.5)</td>
<td>192 (74.7)</td>
<td>60 (69.8)</td>
<td>0.4489</td>
<td>343</td>
</tr>
<tr>
<td>Comorbidities</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hypertension</td>
<td>259 (75.5)</td>
<td>183 (71.2)</td>
<td>76 (88.4)</td>
<td>0.0022</td>
<td>343</td>
</tr>
<tr>
<td>Diabetes mellitus</td>
<td>92 (26.8)</td>
<td>57 (22.2)</td>
<td>35 (40.7)</td>
<td>0.0013</td>
<td>343</td>
</tr>
<tr>
<td>Hyperlipidemia</td>
<td>198 (57.7)</td>
<td>149 (58)</td>
<td>49 (57)</td>
<td>0.9710</td>
<td>343</td>
</tr>
<tr>
<td>Smoking status</td>
<td></td>
<td></td>
<td></td>
<td>0.0265</td>
<td>338</td>
</tr>
<tr>
<td>Current smoker</td>
<td>135 (39.9)</td>
<td>113 (44)</td>
<td>22 (27.2)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>Ex-smoker</td>
<td>99 (29.3)</td>
<td>70 (27.2)</td>
<td>29 (35.8)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>Never smoker</td>
<td>104 (30.8)</td>
<td>74 (28.8)</td>
<td>30 (37)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>Clinical characteristics</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Prehospital time in minutes (median, IQR)</td>
<td>120 (70–273.5)</td>
<td>125 (80–271)</td>
<td>103 (58.5–277.5)</td>
<td>0.0554</td>
<td>324</td>
</tr>
<tr>
<td>Systolic blood pressure at admission (median, IQR)</td>
<td>144 (128–160)</td>
<td>140 (125–160)</td>
<td>148 (130.5–163.25)</td>
<td>0.1722</td>
<td>343</td>
</tr>
<tr>
<td>Diastolic blood pressure at admission (median, IQR)</td>
<td>80 (70–97)</td>
<td>80 (70–100)</td>
<td>80 (70–94)</td>
<td>0.1569</td>
<td>340</td>
</tr>
<tr>
<td>Heart rate at admission (median, IQR)</td>
<td>76 (67–88)</td>
<td>76 (66.75–88)</td>
<td>80 (67–90)</td>
<td>0.1517</td>
<td>341</td>
</tr>
<tr>
<td>Left ventricular EF</td>
<td></td>
<td></td>
<td></td>
<td>0.0042</td>
<td>333</td>
</tr>
<tr>
<td>&gt;50%</td>
<td>163 (48.9)</td>
<td>131 (52.6)</td>
<td>32 (38.1)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>41–50%</td>
<td>75 (22.5)</td>
<td>58 (23.3)</td>
<td>17 (20.2)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>31–40%</td>
<td>74 (22.2)</td>
<td>50 (20.1)</td>
<td>24 (28.6)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>≤30%</td>
<td>21 (6.3)</td>
<td>10 (4)</td>
<td>11 (13.1)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>Kidney function</td>
<td></td>
<td></td>
<td></td>
<td>&lt;0.001</td>
<td>343</td>
</tr>
<tr>
<td>eGFR ≥60 ml/min/1.73 m²</td>
<td>247 (72)</td>
<td>202 (78.6)</td>
<td>45 (52.3)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>eGFR 30–59 ml/min/1.73 m²</td>
<td>85 (24.8)</td>
<td>54 (21)</td>
<td>31 (36)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>eGFR &lt;30 ml/min/1.73 m²</td>
<td>11 (3.2)</td>
<td>1 (0.4)</td>
<td>10 (11.6)</td>
<td>– –</td>
<td></td>
</tr>
<tr>
<td>Treatment</td>
<td></td>
<td></td>
<td></td>
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<td></td>
</tr>
<tr>
<td>PCI</td>
<td>314 (91.5)</td>
<td>240 (93.4)</td>
<td>74 (86)</td>
<td>0.0583</td>
<td>343</td>
</tr>
<tr>
<td>Bypass therapy</td>
<td>34 (9.9)</td>
<td>22 (8.6)</td>
<td>12 (14)</td>
<td>0.2149</td>
<td>343</td>
</tr>
<tr>
<td>Lysis therapy</td>
<td>2 (0.6)</td>
<td>2 (0.8)</td>
<td>0 (0)</td>
<td>0.9981</td>
<td>343</td>
</tr>
<tr>
<td>Laboratory values</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Troponin I at admission (median, IQR)</td>
<td>0.555 (0.08–4.9525)</td>
<td>0.490 (0.0775–4.505)</td>
<td>1.085 (0.0975–6.6525)</td>
<td>0.1388</td>
<td>336</td>
</tr>
<tr>
<td>peak CRP (median, IQR)</td>
<td>0.37 (0.29–0.985)</td>
<td>0.34 (0.29–0.81)</td>
<td>0.54 (0.29–1.55)</td>
<td>0.0081</td>
<td>342</td>
</tr>
</tbody>
</table>
and PCI did not change the association substantially (see Figure 2). In this model, FGF-23 produced an FDR-adjusted p-value of <0.001 and a Hazard Ratio (HR) of 1.57 [95% CI: 1.29–1.91] per SD.

In the subsequently performed 5-years ROC analysis, 339 patients were included (for 1 patient there was no information on 5 year survival). The number of events (=death within 5 years) was 44 and the number of controls was 295. Figure 3A shows the ROC curve for FGF-23, and the AUC was 0.6903 [95% CI: 0.5937–0.781]. On the basis of the ROC analysis we calculated the optimal cut-off value for a dichotomization of the standardized FGF-23 variable, which was at 0.4186204 SD (Youden-Index: 1.357627). Using this cut-off for categorization of the full study sample, 64 patients were assigned to the "high FGF-23 group" and 276 patients were assigned to the "low FGF-23 group". Figure 3B displays the Kaplan-Meier curves of those two groups (high/low FGF-23 group).

4. Discussion

In this study, we investigated the association of 92 plasma proteins with long-term mortality in ST-elevation myocardial infarction patients. One protein, FGF-23, exhibited a particularly strong association with long-term mortality in the multivariable-adjusted Cox regression model (HR: 1.57, 95% CI: 1.29–1.91) as well as in a 5 years ROC analysis (AUC: 0.6903, 95% CI: 0.5937–0.781). All other plasma proteins were not significantly related to long-term mortality in the multivariable-adjusted Cox regression model.

FGF-23 is a representative of the fibroblast growth factor family and is classified by OLINK as a cancer marker. It is mainly produced by osteoblasts and osteocytes but can also be produced and secreted by other tissue cells (11). Several prior studies reported associations between FGF-23 and mortality in the general population (12–15). There is also strong evidence for an association between elevated FGF-23 concentrations and cardiovascular events in general and mortality in coronary artery disease (CAD) and after AMI in particular (16–20). Nevertheless, the majority of studies on FGF-23 and mortality in CAD patients and after AMI used short- to mid-term mortality as the primary endpoint and so it remains unclear, whether FGF-23 concentrations are also associated with long-term mortality after AMI and whether the effects remain stable over the course of several years. With a median follow-up time of 7.6 years, this study was able to address this question indicating a long-lasting association up to one decade after the acute event.

4.1. Causality of the relationship

In a meta-analysis on FGF-23 and risks of cardiovascular and non-cardiovascular diseases the authors concluded that the association between FGF-23 and the diseases might be non-
causal (21). In a rat model, Andrukhova et al. found that the induction of myocardial infarction in rats led to an elevation of FGF-23 concentrations (22), which strongly indicates a principal relation between AMI and FGF-23. Likewise, Schumacher et al. reported that in mice cardiac fibroblasts located in infarcted myocardial tissue produced and expressed FGF-23 in response to

![Graph showing ROC curve and Kaplan-Meier plot](image-url)
the induced AMI (23). Thus, results of these studies suggest that AMI causes an elevation of plasma FGF-23 concentrations and do not support the idea of elevated FGF-23 concentrations being a risk factor for AMI and AMI-related mortality. In a further study, Hao et al. investigated the role of FGF-23 in post-infarct myocardial fibrosis in mice. Their findings indicate that FGF-23 actually promotes myocardial fibrosis and exacerbates diastolic dysfunction induced by MI. To sum up, the scientific data on the causal nature of the relationship between FGF-23 and AMI is not very clear at the moment.

### 4.2. FGF-23 and left ventricular (LV) remodeling

It has been reported that high plasma FGF-23 concentrations are independently associated with LV remodeling after STEMI (24). Pathological ventricular remodeling is a major complication after AMI and affects cardiac function and overall prognosis in the long-term (25). Likely, pathological ventricular remodeling plays a major role in the association between elevated FGF-23 plasma concentrations and long-term mortality in AMI patients. Nevertheless, in our study no data on ventricular remodeling over the course of the years was available and therefore we are not able to further examine these supposed interconnections.

### 4.3. FGF-23 and chronic kidney disease (CKD)

Other pathophysiological pathways might potentially be involved in the association between plasma FGF-23 and long-term outcome after STEMI. First, FGF-23 is known to be strongly connected to CKD, which is a major risk factor for adverse outcomes in AMI patients (26–29). Consequently, it could be speculated that the predictive ability of FGF-23 for long-term outcome in STEMI patients might substantially be driven by the association between FGF-23 and kidney function. Nevertheless, the multivariable Cox regression model was adjusted for the numerical variable GFR (ml/min/1.73 m²), so we conclude that the association between FGF-23 and mortality after AMI is (mainly) independent of renal function or CKD.

### 4.4. Heart failure and reduced left-ventricular ejection fraction (LVEF)

The scientific literature moreover suggests a relation between FGF-23 and heart failure/reduced left-ventricular ejection fraction (30–32). FGF-23 is also a strong predictor of outcome in heart failure patients (33–35). It is well known, that impaired LVEF goes along with higher long-term mortality in AMI patients (36, 37). In a study by Cornelisens et al., FGF-23 was only predictive for 1-year mortality in AMI patients with heart failure, but not so in AMI patients without heart failure; which indicates an interaction with heart failure and LVEF. Thus, we tested this hypothesis but couldn’t find a significant interaction between LVEF and FGF-23 plasma concentrations in our data. We moreover hypothesized, that the association between elevated FGF-23 levels and long-term mortality after AMI could partially be mediated by a reduced LVEF. So we calculated a multivariable Cox regression model including LVEF (see Supplementary Figure S1), but again, there was almost no change to the original model.

### 4.5. Prehospital time and therapeutic delay

The blood samples in this study were very uniformly taken at the beginning of the cardiac catheterization. As we included only STEMI events, the delay between hospital admission and PCI was generally very small. Nevertheless, greater deviations between the patients were observed in prehospital time, which represent the majority of therapeutic delay and negatively affects outcome in AMI patients (38, 39). Takahashi et al. found that serum FGF-23 concentrations after AMI treated with PCI follow a certain course of time after AMI with a slight decrease on day 1 and 3 after admission and an increase on day 5 and 7 (40); Thorsen et al. came to similar results reporting a reduction in FGF-23 concentrations during the acute phase of AMI and a normalization after seven days following successful revascularization (41). To take a closer look at this and in order to examine whether the association between FGF-23 and mortality was mainly driven by the therapeutic delay, we calculated another multivariable adjusted model including prehospital time in minutes (see Supplementary Figure S2). Although results of the model were influenced to a limited extent, it shows that therapeutic delay does not play a superior role in the association between FGF-23 and long-term mortality.

### 4.6. Strengths and limitation

This study is characterized by some strengths. First, it is based on patients from the population-based Augsburg Myocardial Infarction Registry with consecutive enrollment of patients, which minimizes the effect of selection bias. Blood samples were uniformly taken during the PTCA intervention, guaranteeing highest consistency in blood sampling. Moreover, for every case there was a large number of additional information which we used for proper adjustment in the Cox regression models.

Nevertheless, there are also some limitations. We have no validation cohort from another register for our analyses, which means we cannot validate the associations found in this study. Moreover, as this study is based solely on observational data, we cannot draw any conclusions about causality (including the possibility of reverse causality). Additionally, we might not have considered all potential confounders. Last, as this study included only patients between 25 and 84 years with STEMI, the results may not be generalized to all age or ethnic groups as well as to Non-ST-elevation MI events.
5. Conclusion

Elevated FGF-23 plasma concentrations are independently associated with higher long-term mortality in STEMI patients. The effects appear to remain stable over the course of one decade. FGF-23-concentrations at baseline might help to identify a population in the need of intensive and lasting secondary prevention measures to target long-term mortality. Moreover, if a causal nature of the relationship can be established, new therapies targeting FGF-23 might represent a promising approach especially in the long-term treatment of patients after AMI.

Data availability statement

The datasets generated during and/or analyzed in the current study are not publicly available due to data protection aspects but are available in an anonymized form from the corresponding author on reasonable request. Requests to access these datasets should be directed to TS, timo.schmitz@med.uni-augsburg.de.

Ethics statement

The studies involving human participants were reviewed and approved by the Ethics Committee of the Bavarian Medical Association. The patients/participants provided their written informed consent to participate in this study.

Author contributions

TS and JL conceived the study. TS performed the statistical analysis and drafted the manuscript. JL supervised data analysis. CM was responsible for the acquisition of the data. BW, MH, AP, and JL contributed to data acquisition and revised the manuscript. All authors contributed to the article and approved the submitted version.

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Conflict of interest

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Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fcvm.2023.1173281/full#supplementary-material

References


