



# Short-term exposure to ambient nitrogen dioxide and fine particulate matter and cause-specific mortality: A causal modeling approach in four regions



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## ABSTRACT

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Ambient air pollution still represents a major health burden. While the link between short-term air pollution exposures and mortality has been well-documented globally, few studies have applied causal modeling approaches. Therefore, we aimed to quantify the relationship between day-to-day changes in ambient particulate matter with an aerodynamic diameter  $\leq 2.5 \mu\text{m}$  ( $\text{PM}_{2.5}$ ) and nitrogen dioxide ( $\text{NO}_2$ ) levels and changes in daily natural, cardiovascular (including all-cardiovascular, cardiac, and stroke), as well as respiratory mortality rates using a causal modeling framework. Daily air pollution data and cause-specific death counts at the county, district, or municipality level from California (US), Jiangsu (China), Germany, and Lazio (Italy) were obtained for the years 2015–2019, including urban and rural populations. We used interactive fixed effects models to analyze the effects of air pollutants across different lag periods (0–2, 3–7, and 0–7 days after exposure) while accounting for both measured and unmeasured time-varying spatial unit-specific confounding factors. We observed increases in daily cardiovascular deaths (per 1 million people) per a  $10 \mu\text{g}/\text{m}^3$  increase in daily  $\text{NO}_2$  at lag 0–7: 0.18 (95 % confidence interval: 0.02, 0.38) in California, 0.23 (0.14, 0.32) in Jiangsu, 0.48 (0.27, 0.70) in Germany, and  $-0.35$  ( $-2.63$ , 1.92) in Lazio. For  $\text{PM}_{2.5}$ , the related increases in cardiovascular mortality rates were 0.00 ( $-0.18$ , 0.18) in California, 0.04 (0.00, 0.09) in Jiangsu, 0.22 (0.06, 0.37) in Germany, and 1.96 (0.76, 3.16) in Lazio. Additionally, associations were seen for natural, cardiac, stroke, and respiratory mortality, particularly pronounced among individuals aged 75 and older. These associations were strongest with prolonged exposures and remained consistent even in two-pollutant models. This study, using a causal modeling approach and including urban and rural populations, contributes to the growing body of evidence linking increases in short-term exposure to  $\text{NO}_2$  and  $\text{PM}_{2.5}$  with increased cause-specific mortality rates.

## 1. Introduction

Ambient air pollution is a major health risk, causing 4.5 million premature deaths annually globally (GBD 2019 Risk Factors Collaborators, 2020). The World Health Organization (WHO) rates short-term exposure to particulate matter (PM) with an aerodynamic diameter  $\leq 2.5 \mu\text{m}$  ( $\text{PM}_{2.5}$ ) as causally linked to all-cause, cardiovascular, and

respiratory mortality, while the effect of nitrogen dioxide ( $\text{NO}_2$ ) on all-cause mortality is suggestive (WHO, 2021), but supported by recent systematic reviews confirming positive associations (Atkinson et al., 2014; Orellano et al., 2020; Wang et al., 2021).

Although initially affecting the respiratory system, air pollutants pose the highest attributable risk to the cardiovascular system (Brook et al., 2010), as shown in numerous epidemiological studies worldwide,

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even at low levels of exposure (Newby et al., 2015; Schwartz et al., 2017; Thurston et al., 2017). While the short-term effects of PM<sub>2.5</sub> on cardiovascular health are well established (Brook et al., 2010), the independent contribution of NO<sub>2</sub> is still widely debated (U.S. EPA, 2016).

Previous studies of the acute effects of NO<sub>2</sub> or PM<sub>2.5</sub> on cause-specific mortality have assessed metropolitan areas of the same country (Linares et al., 2018; Chen et al., 2019), multiple cities across the world (Liu et al., 2019; Meng et al., 2021), or single geographic regions (Stafoggia et al., 2020; Liu et al., 2021; Liu et al., 2022; Gariazzo et al., 2023). This focus might stem from a lack of air quality data in rural settings or sparse daily cause-specific mortality data with fine spatial resolution. However, epidemiological findings may lack generalizability because of potential disparities in population characteristics and source profiles of air pollutants in rural and urban areas, especially at very low concentrations (Gariazzo et al., 2023). Geographic differences in the association between NO<sub>2</sub> or PM<sub>2.5</sub> and cause-specific mortality emphasize the need for further research (Atkinson et al., 2014; Wang et al., 2021).

Most epidemiological studies have used time-series analyses to examine short-term effects (Atkinson et al., 2014; Orellano et al., 2020; Wang et al., 2021). However, these might be biased by unmeasured temporal confounding (Schwartz et al., 2017; Wei et al., 2020; Guo et al., 2023), suggesting that causal modeling approaches are needed.

The objective of the current study was to examine the association between daily changes in air pollution levels and day-to-day changes in cause-specific mortality rates in four different regions on three continents, including urban and rural populations: California, United States (US); Jiangsu Province, China (hereinafter Jiangsu); Germany; and Lazio Region, Italy (hereinafter Lazio). We used so-called interactive fixed effects (IFE) models (Ma et al., 2024) – a more flexible generalization of two-way fixed effects models – to investigate the association between short-term exposures to NO<sub>2</sub> and PM<sub>2.5</sub> and natural, all-cardiovascular, cardiac, stroke, as well as respiratory mortality rates. As a further objective, we explored potential variations in these effects based on sex, age, and urbanicity.

## 2. Methods

We used anonymized daily county-, district-, or municipality-level mortality records. This study received approval from the Yale Institutional Review Boards (IRB protocol ID: 2000029741) and adheres to the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) reporting guidelines (von Elm et al., 2007).

### 2.1. Mortality data

Daily cause-specific death counts by sex and age from 2015 to 2019 were collected in each spatial unit (county in California and Jiangsu; district in Germany; and municipality in Lazio) from national, state, or regional registries (eTable 1) based on each country's finest available administrative unit of mortality data. In Lazio, we limited our analysis to 79 (out of 378) municipalities, each with  $\geq 10,000$  population, to ensure an adequate number of deaths.

To define the mortality outcomes, we used both the 9th Revision (for Lazio, 2015–2018 only) and the 10th Revisions of the International Classification of Diseases. The outcomes include: Natural (i.e., non-accidental) (ICD-9: 1–799; ICD-10: A00–R99), cardiovascular (390–459; I00–I99), cardiac (390–429; I00–I52), cerebrovascular (430–438; I60–I69), and respiratory mortality (460–519; J00–J99). Using corresponding annual population size data (eTable 2), we computed daily cause-specific mortality rates at the spatial unit level for the entire population and population subgroups by sex (male and female) and age (0–74 and  $\geq 75$  years).

### 2.2. Environmental data

Daily mean concentrations of NO<sub>2</sub> and PM<sub>2.5</sub> were available from

spatiotemporal models in 1 km  $\times$  1 km grid cells for Jiangsu (Huang et al., 2021; Huang et al., 2022), Germany (Flemming and Stern, 2004; Nordmann et al., 2020), and Lazio (Stafoggia et al., 2019; Stafoggia and Bellander, 2020), as well as data from air quality monitoring sites for California (U.S. EPA, 2020), including data from 32 (out of 58) counties with monitoring sites for both NO<sub>2</sub> and PM<sub>2.5</sub>. For California, data from air quality monitoring sites were used since, at the time of our study, no high-resolution daily spatiotemporal data for NO<sub>2</sub> and PM<sub>2.5</sub> covering the whole of California were publicly available covering 2015 to 2019. For Jiangsu, Germany, and Lazio, we estimated daily mean pollutant concentrations in each spatial unit by calculating area-weighted averages within grids that intersected with respective spatial units. In California counties with multiple monitoring sites, average air pollution concentrations were calculated at the county level (eMethods 1).

We obtained hourly air temperature and dewpoint temperature data at a resolution of  $0.1^\circ \times 0.1^\circ$  (9 km  $\times$  9 km) from the ERA5-Land reanalysis dataset for all regions (Muñoz Sabater, 2019). Similar to the air pollution data, daily air temperature and dewpoint temperature averages and corresponding lags were computed for each spatial unit. Based on these data, we also calculated relative humidity and apparent temperature (eMethods 2).

### 2.3. Statistical analysis

We calculated descriptive statistics for air pollutants, air temperature, and cause-specific mortality rates. Spearman correlation coefficients were used to evaluate temporal variations.

Associations between air pollutants and cause-specific mortality were examined using a novel IFE model, a causal modeling approach. This model, detailed in Ma et al., controls for confounders varying across space (but not in time), for confounders varying across time (but not in space), and also for unmeasured time-varying spatial unit-specific confounders (Ma et al., 2024).

The outcome of our IFE model is the day-to-day difference in the spatial unit cause-specific mortality rate; the exposure is the day-to-day difference in the spatial unit average concentration of NO<sub>2</sub> or PM<sub>2.5</sub>. The day-to-day change in air temperature on the same lag day as the air pollution variable is included as a natural cubic spline with five degrees of freedom (df). Day-to-day variations (i.e., first-order differences) were calculated to remove the long-term and seasonal trends of mortality rate, air pollution, and air temperature and to meet the stationarity and normality assumptions of the model. Further, the model includes a variable for time-invariant spatial unit effects and a variable for unmeasured time-varying spatial unit effects. More details on the IFE model are provided in the eMethods 3. The results of the IFE models are presented as expected changes in the day-to-day cause-specific mortality rate for each unit change in the daily air pollution level.

The short-term associations between changes in daily mortality rates and changes in exposure to air pollution were explored focusing on immediate (i.e., 0–2 days after exposure [lag0–2]), delayed [lag3–7], and prolonged [lag0–7] effects. In addition to single-pollutant models, we also conducted two-pollutant models, incorporating both NO<sub>2</sub> and PM<sub>2.5</sub> in the same model, to address possible mutual confounding and estimate independent effects. A univariate Cochran Q-test was performed to examine the heterogeneity of estimates across regions, and a random effects meta-analytical approach using restricted maximum likelihood (REML) was applied to calculate pooled estimates (Sera et al., 2019).

To identify potentially sensitive subpopulations, we examined effect modification by analyzing sex- (male and female) and age-stratified (0–74 years and  $\geq 75$  years) data. In addition, we performed stratified analyses to explore potential effect modification by urbanicity (urban and rural) as detailed in eTable 3. We calculated z-scores to test whether there were significant differences in air pollutant effects between the groups (Ma et al., 2023).

We performed several sensitivity analyses to test the robustness of our results (e.g., different model parameters or confounding variables)

and also explored effects on cardiovascular mortality at each single lag day (eMethods 4).

All analyses were conducted using R software, version 4.1.3 (R Core Team, 2023). The IFE analyses were performed by applying the package *phtt* (Bada and Liebl, 2014).

A two-sided P-value  $<0.05$  was considered statistically significant.

### 3. Results

#### 3.1. Descriptive analysis

Overall, this analysis comprised 7,414,152 deaths from natural causes, including 2,904,720 and 721,325 deaths from cardiovascular and respiratory diseases, respectively, in California, Jiangsu, Germany, and Lazio from 2015 to 2019. Total numbers of cause-specific deaths by sex, age group, and urbanicity are presented in Table 1. Table 2 shows descriptive statistics of each region's daily cause-specific mortality rates, air pollution, and air temperature. The description of sex- and age-specific mortality rates are presented in eTable 4 and eTable 5. Descriptive statistics of mortality rates and environmental factors stratified by urbanicity are displayed in eTable 6. Fig. 1 illustrates the spatial distribution of NO<sub>2</sub> and PM<sub>2.5</sub> concentrations across the regions. The correlation among air pollutants and meteorological factors is displayed in eTable 7. There were no missing data in this study.

#### 3.2. Association of air pollution with cause-specific mortality

The estimated associations between day-to-day changes in NO<sub>2</sub> and

PM<sub>2.5</sub> levels and cause-specific mortality rates (per 1 million people) at different lags (0–2, 3–7, 0–7) from both single- and two-pollutant models are displayed in Fig. 2 and eTable 8. In single-pollutant models, increases in NO<sub>2</sub> or PM<sub>2.5</sub> were significantly associated with increased mortality rates across lags and different causes of death. Generally, estimates were strongest for natural mortality, showed considerably wider confidence intervals (CI) in Lazio than in the other three regions, and were most pronounced in prolonged lags (i.e., the 8-day average). Therefore, the estimates of lag0-7 were used as the main exposure in further analyses.

For example, we observed a relatively stable lag pattern across the four study regions for the effects of daily changes in air pollutants on daily changes in cardiovascular mortality rate. In single-pollutant models, the highest estimates were for lag0-7 for NO<sub>2</sub> in Germany (0.48 (95 % CI: 0.27, 0.70)) and PM<sub>2.5</sub> in Lazio (1.96 (95 % CI: 0.76, 3.16)). Cardiac mortality, as a subcategory of cardiovascular mortality, showed similar effect estimates and lag patterns, although the effects were commonly smaller. Changes in cerebrovascular mortality rates were more pronounced at the immediate lag0-2 and the prolonged lag0-7.

The lag pattern of the association between daily changes in air pollution and daily changes in respiratory mortality was inconsistent across study regions. It showed the highest effect estimates at lag0-7 for NO<sub>2</sub> in Germany (0.24 (95 % CI: 0.15, 0.34)) and PM<sub>2.5</sub> in Lazio (0.58 (95 % CI: 0.01, 1.15)).

The results from two-pollutant models were generally consistent with those from single-pollutant models, though the estimated coefficients for NO<sub>2</sub> in Lazio and PM<sub>2.5</sub> in Jiangsu were relatively smaller.

**Table 1**  
Cause-specific deaths during the period 2015–2019 in the four study regions.

|                       | Cause: | Natural   | Cardiovascular | Cardiac   | Cerebrovascular | Respiratory |
|-----------------------|--------|-----------|----------------|-----------|-----------------|-------------|
| <i>California, US</i> |        |           |                |           |                 |             |
| Total                 |        | 1,055,977 | 373,760        | 293,788   | 68,222          | 103,837     |
| Sex                   | Male   | 525,775   | 191,080        | 156,333   | 28,836          | 50,617      |
|                       | Female | 530,176   | 182,678        | 137,453   | 39,386          | 53,220      |
| Age-group             | 0–74   | 391,830   | 113,870        | 92,794    | 16,712          | 30,012      |
|                       | 75+    | 664,077   | 259,869        | 200,977   | 51,507          | 73,817      |
| Urbanicity            | Urban  | 827,210   | 295,886        | 232,486   | 53,901          | 79,688      |
|                       | Rural  | 228,767   | 77,874         | 61,302    | 14,321          | 24,149      |
| <i>Jiangsu, China</i> |        |           |                |           |                 |             |
| Total                 |        | 2,425,063 | 993,171        | 412,991   | 574,564         | 320,520     |
| Sex                   | Male   | 1,327,688 | 497,304        | 201,556   | 292,371         | 178,480     |
|                       | Female | 1,097,336 | 495,843        | 211,424   | 282,180         | 142,036     |
| Age-group             | 0–74   | 930,036   | 290,190        | 117,456   | 170,210         | 66,398      |
|                       | 75+    | 1,494,919 | 702,981        | 295,535   | 404,354         | 254,116     |
| Urbanicity            | Urban  | 1,192,833 | 467,781        | 191,700   | 273,087         | 142,288     |
|                       | Rural  | 1,232,230 | 525,390        | 221,291   | 301,477         | 178,232     |
| <i>Germany</i>        |        |           |                |           |                 |             |
| Total                 |        | 3,712,220 | 1,455,883      | 1,172,858 | 220,813         | 278,713     |
| Sex                   | Male   | 1,764,397 | 636,519        | 518,366   | 89,771          | 146,269     |
|                       | Female | 1,947,823 | 819,364        | 654,492   | 131,042         | 132,444     |
| Age-group             | 0–74   | 995,525   | 239,678        | 191,405   | 33,818          | 71,685      |
|                       | 75+    | 2,716,695 | 1,216,205      | 981,453   | 186,995         | 207,028     |
| Urbanicity            | Urban  | 2,472,450 | 922,206        | 735,300   | 145,373         | 189,989     |
|                       | Rural  | 1,239,770 | 533,677        | 437,558   | 75,440          | 88,724      |
| <i>Lazio, Italy</i>   |        |           |                |           |                 |             |
| Total                 |        | 220,892   | 81,906         | 61,959    | 16,922          | 18,255      |
| Sex                   | Male   | 104,444   | 35,874         | 27,515    | 6,745           | 8,881       |
|                       | Female | 116,448   | 46,032         | 34,444    | 10,177          | 9,374       |
| Age-group             | 0–74   | 55,945    | 13,317         | 10,167    | 2,327           | 2,624       |
|                       | 75+    | 164,947   | 68,589         | 51,792    | 14,595          | 15,631      |
| Urbanicity            | Urban  | 137,468   | 50,184         | 38,215    | 10,117          | 11,449      |
|                       | Rural  | 83,424    | 31,722         | 23,744    | 6,805           | 6,806       |
| <i>All regions</i>    |        |           |                |           |                 |             |
| Total                 |        | 7,414,152 | 2,904,720      | 1,941,596 | 880,521         | 721,325     |
| Sex                   | Male   | 3,722,304 | 1,360,777      | 903,770   | 417,723         | 384,247     |
|                       | Female | 3,691,783 | 1,543,917      | 1,037,813 | 462,785         | 337,074     |
| Age-group             | 0–74   | 2,373,336 | 657,055        | 411,822   | 223,067         | 170,719     |
|                       | 75+    | 5,040,638 | 2,247,644      | 1,529,757 | 657,451         | 550,592     |
| Urbanicity            | Urban  | 4,629,961 | 1,736,057      | 1,197,701 | 482,478         | 423,414     |
|                       | Rural  | 2,784,191 | 1,168,663      | 743,895   | 398,043         | 297,911     |

**Table 2**

Daily county-/district-/municipality-level cause-specific mortality rates and environmental data, 2015–2019, in the four study regions.

|   | Mean (SD)   | 25th Percentile | Median (IQR) | 75th Percentile |
|---|-------------|-----------------|--------------|-----------------|
| <b>California, US (32 counties, total population 37,290,255, total area 247,411 km<sup>2</sup>)<sup>a</sup></b> |             |                 |              |                 |
| <i>Cause-specific mortality rate (per 1 million people)</i>   |             |                 |              |                 |
| Natural   | 15.6 (7.2)  | 11.5            | 15.1 (7.6)   | 19.1            |
| Cardiovascular  | 5.3 (3.9)   | 2.8             | 5.1 (4.2)    | 7.1             |
| Cardiac   | 4.1 (3.4)   | 2.0             | 3.9 (3.6)    | 5.7             |
| Cerebrovascular   | 1.0 (1.7)   | 0.0             | 0.0 (1.4)    | 1.4             |
| Respiratory   | 1.6 (2.1)   | 0.0             | 1.1 (2.3)    | 2.3             |
| <i>Environmental factors</i>  |             |                 |              |                 |
| NO <sub>2</sub> (µg/m <sup>3</sup> )  | 15.7 (11.4) | 7.5             | 12.8 (13.6)  | 21.1            |
| Urban   | 19.6 (13.1) | 9.5             | 16.3 (16.9)  | 26.4            |
| Rural   | 12.7 (8.9)  | 6.2             | 10.8 (10.9)  | 17.1            |
| PM <sub>2.5</sub> (µg/m <sup>3</sup> )  | 9.9 (9.2)   | 5.3             | 8.0 (6.4)    | 11.7            |
| Urban   | 9.3 (7.6)   | 5.4             | 8.0 (6.0)    | 11.4            |
| Rural   | 10.4 (10.3) | 5.3             | 8.1 (6.8)    | 12.1            |
| Air temperature (°C)  | 16.3 (7.0)  | 11.2            | 15.7 (10.1)  | 21.2            |
| <b>Jiangsu, China (82 counties, total population 77,955,026, total area 102,010 km<sup>2</sup>)</b>             |             |                 |              |                 |
| <i>Cause-specific mortality rate (per 1 million people)</i>   |             |                 |              |                 |
| Natural   | 17.2 (6.9)  | 12.4            | 16.4 (8.7)   | 21.1            |
| Cardiovascular  | 6.9 (3.9)   | 4.2             | 6.4 (4.8)    | 9.0             |
| Cardiac   | 2.8 (2.3)   | 1.2             | 2.5 (2.8)    | 4.0             |
| Cerebrovascular   | 4.0 (2.7)   | 2.1             | 3.7 (3.4)    | 5.5             |
| Respiratory   | 2.2 (2.2)   | 0.5             | 1.8 (2.9)    | 3.3             |
| <i>Environmental factors</i>  |             |                 |              |                 |
| NO <sub>2</sub> (µg/m <sup>3</sup> )  | 32.2 (12.8) | 23.1            | 29.8 (15.1)  | 38.2            |
| Urban   | 36.2 (14.3) | 26.0            | 33.2 (17.1)  | 43.1            |
| Rural   | 28.2 (9.4)  | 21.1            | 26.7 (12.7)  | 33.8            |
| PM <sub>2.5</sub> (µg/m <sup>3</sup> )  | 50.8 (27.8) | 31.0            | 44.0 (31.8)  | 62.8            |
| Urban   | 49.6 (28.0) | 29.9            | 42.8 (31.7)  | 61.6            |
| Rural   | 52.0 (27.5) | 32.0            | 45.3 (32.0)  | 64.0            |
| Air temperature (°C)  | 16.1 (9.1)  | 8.0             | 17.1 (15.7)  | 23.8            |
| <b>Germany (401 districts, total population 82,735,005, total area 357,672 km<sup>2</sup>)</b>                  |             |                 |              |                 |
| <i>Cause-specific mortality rate (per 1 million people)</i>   |             |                 |              |                 |
| Natural   | 25.0 (15.4) | 14.8            | 23.3 (18.2)  | 33.0            |
| Cardiovascular  | 10.2 (9.6)  | 2.2             | 8.6 (12.7)   | 15.0            |
| Cardiac   | 8.3 (8.6)   | 0.0             | 7.0 (12.3)   | 12.3            |
| Cerebrovascular   | 1.5 (3.5)   | 0.0             | 0.0 (0.0)    | 0.0             |
| Respiratory   | 1.9 (4.0)   | 0.0             | 0.0 (2.1)    | 2.1             |
| <i>Environmental factors</i>  |             |                 |              |                 |
| NO <sub>2</sub> (µg/m <sup>3</sup> )  | 12.2 (7.8)  | 6.7             | 10.1 (8.8)   | 15.5            |
| Urban   | 15.0 (8.9)  | 8.5             | 12.7 (10.8)  | 19.3            |
| Rural   | 9.5 (5.4)   | 5.7             | 8.2 (6.3)    | 12.0            |
| PM <sub>2.5</sub> (µg/m <sup>3</sup> )  | 10.0 (7.0)  | 5.5             | 8.2 (6.8)    | 12.3            |
| Urban   | 10.2 (7.1)  | 5.7             | 8.5 (6.9)    | 12.6            |
| Rural   | 9.8 (7.0)   | 5.3             | 8.0 (6.7)    | 12.1            |
| Air temperature (°C)  | 10.2 (7.4)  | 4.3             | 10.0 (11.9)  | 16.2            |
| <b>Lazio, Italy (79 municipalities, total population 5,104,384, total area 7268 km<sup>2</sup>)<sup>b</sup></b> |             |                 |              |                 |
| <i>Cause-specific mortality rate (per 1 million people)</i>   |             |                 |              |                 |
| Natural   | 22.4 (34.1) | 0.0             | 0.0 (40.2)   | 40.2            |
| Cardiovascular  | 8.7 (21.3)  | 0.0             | 0.0 (0.0)    | 0.0             |
| Cardiac   | 6.5 (18.4)  | 0.0             | 0.0 (0.0)    | 0.0             |
| Cerebrovascular   | 1.9 (9.8)   | 0.0             | 0.0 (0.0)    | 0.0             |
| Respiratory   | 1.8 (9.4)   | 0.0             | 0.0 (0.0)    | 0.0             |
| <i>Environmental factors</i>  |             |                 |              |                 |
| NO <sub>2</sub> (µg/m <sup>3</sup> )  | 13.2 (6.3)  | 8.7             | 11.6 (7.5)   | 16.2            |
| Urban   | 16.4 (7.1)  | 10.9            | 14.7 (9.5)   | 20.3            |
| Rural   | 13.1 (6.2)  | 8.7             | 11.5 (7.4)   | 16.1            |
| PM <sub>2.5</sub> (µg/m <sup>3</sup> )  | 13.3 (6.7)  | 9.2             | 11.8 (5.9)   | 15.0            |
| Urban   | 13.3 (5.6)  | 9.7             | 12.1 (5.4)   | 15.2            |
| Rural   | 13.3 (6.7)  | 9.1             | 11.7 (5.9)   | 15.0            |
| Air temperature (°C)  | 15.3 (6.9)  | 10.1            | 14.9 (10.8)  | 20.9            |

IQR: Interquartile range; NO<sub>2</sub>: Nitrogen dioxide; PM<sub>2.5</sub>: Particulate matter with an aerodynamic diameter  $\leq 2.5 \mu\text{m}$ ; SD: Standard deviation.

<sup>a</sup> In California, only 32 (out of 58) counties with EPA air quality monitoring stations were included in the analysis.

<sup>b</sup> In Lazio, we only included 79 (out of 378) municipalities with a population  $\geq 10,000$ .

We observed considerable heterogeneity of estimates across regions ranging from  $I^2 = 0.00\%$  to  $I^2 = 90.54\%$  (eTable 9). Therefore, pooled estimates should be interpreted with caution for some models (eTable 10).

### 3.3. Effect modification by sex, age, and urbanicity

We used lag0-7 to examine potential effect modifications by sex, age, and urbanicity. In subgroup analyses by sex, no consistent trend across study regions was observed, with few significant differences in effect estimates between males and females (Fig. 3 and eTable 11).

The results of the analyses by age group clearly showed stronger effects among people  $\geq 75$  years for NO<sub>2</sub> and PM<sub>2.5</sub>, respectively, across all study regions (Fig. 3 and eTable 12). For example, in a single-pollutant model for Germany, the estimated change in the daily cardiovascular mortality rate (per 1 million people) in older people was 2.93 (95 % CI 1.24, 4.62) and 1.84 (95 % CI: 0.62, 3.06) per 10 µg/m<sup>3</sup> increases in NO<sub>2</sub> and PM<sub>2.5</sub>, respectively, but only 0.17 (95 % CI 0.08, 0.27) and 0.00 (95 % CI: 0.07, 0.07) in the younger age group. In both single- and two-pollutant models, apart from Lazio, the effect modification by age was more pronounced for NO<sub>2</sub> than for PM<sub>2.5</sub>. Consistent significant effect modification was observed for NO<sub>2</sub> for natural, cardiovascular, cerebrovascular, and respiratory mortality in Jiangsu and Germany, and for PM<sub>2.5</sub> for natural and cardiovascular mortality in Germany and Lazio.

In stratified analyses by urbanicity, we found no consistent patterns of effect modification between urban and rural areas (Fig. 3 and eTable 13).

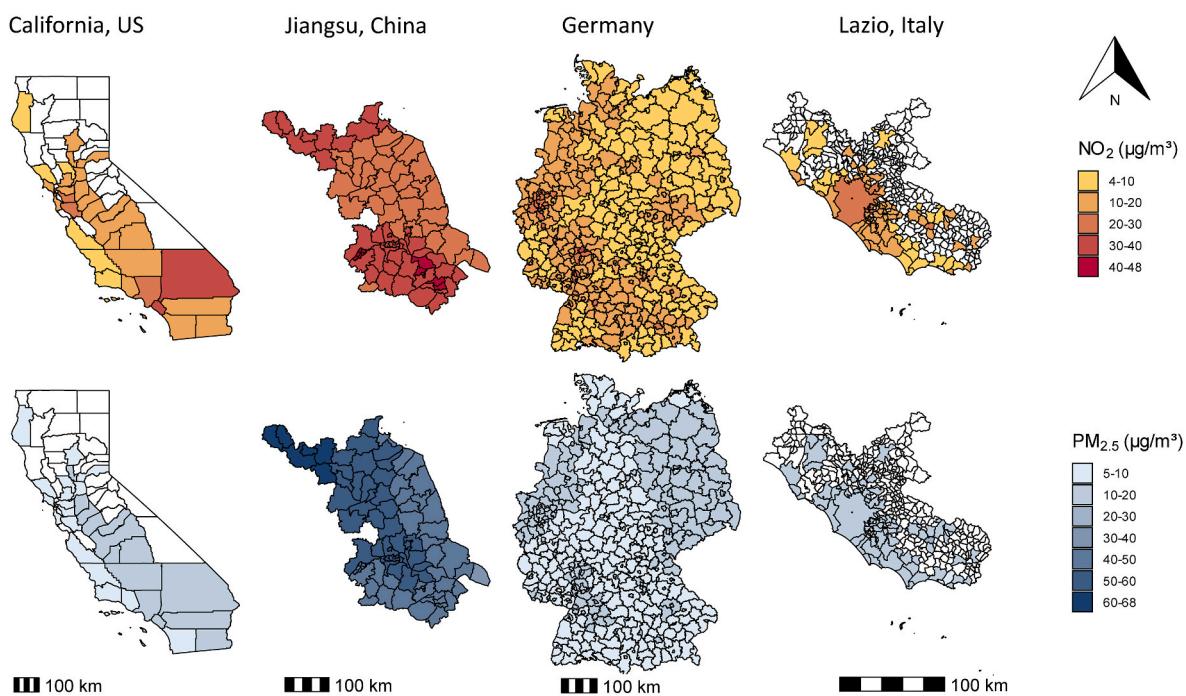
### 3.4. Sensitivity analyses

The results of the sensitivity analyses for cardiovascular mortality are displayed in eFig. 1 and summarized for all mortality outcomes in eTable 14. Our results remained robust in single-pollutant models when changing several model parameters or adjusting for additional variables. In two-pollutant models, the results were robust after restricting the analysis to counties or municipalities with a Spearman correlation coefficient smaller than 0.6 between the two pollutants. The effects of NO<sub>2</sub> and PM<sub>2.5</sub> on cardiovascular mortality at every single lag (0–7) from both single- and two-pollutant models are presented in eTable 15.

## 4. Discussion

In this observational study, we used the IFE model to estimate the short-term associations between daily changes in NO<sub>2</sub> or PM<sub>2.5</sub> and changes in daily cause-specific mortality rates for California, Jiangsu, Germany, and Lazio between 2015 and 2019. We found significant associations between increases in each air pollutant and cause-specific mortality rates across all study regions. Associations were most consistent for natural and cardiovascular mortality, showing pronounced prolonged effects (i.e., lag0-7). The estimated effects were comparable for men and women and urban and rural areas but larger for older people. Our results remained robust in two-pollutant models.

Our findings are generally consistent with the previous literature. For example, in studies of the Multi-Country Multi-City (MCC) Collaborative Research Network, a 10 µg/m<sup>3</sup> increase in NO<sub>2</sub> concentration at lag1 was associated with a 0.46 %, 0.37 %, and 0.47 % increase in all-cause, cardiovascular, and respiratory mortality, respectively (Meng et al., 2021). For PM<sub>2.5</sub>, a 10 µg/m<sup>3</sup> increase in the two-day average



**Fig. 1.** Daily mean NO<sub>2</sub> and PM<sub>2.5</sub> concentrations in each spatial unit.

These maps display the daily average NO<sub>2</sub> and PM<sub>2.5</sub> concentrations in each county, district or municipality in California, US; Jiangsu, China; Germany; and Lazio, Italy (2015–2019). The white areas in California are counties that lacked air quality monitoring stations for both NO<sub>2</sub> and PM<sub>2.5</sub>; the white areas in Lazio represent municipalities with a population of less than 10,000.

NO<sub>2</sub>: Nitrogen dioxide; PM<sub>2.5</sub>: Particulate matter with an aerodynamic diameter  $\leq 2.5 \mu\text{m}$ .

(lag0-1) was associated with an increase of 0.44 %, 0.36 %, and 0.47 % (Liu et al., 2019). Our study's estimates cannot be directly compared to those from time-series or case-crossover studies due to differing interpretations of model coefficients. However, Ma et al. found similar results when comparing the IFE model to a traditional time-series model using the same data. Their findings indicate that short-term exposure to NO<sub>2</sub> or PM<sub>2.5</sub> was linked to increased all-cause mortality (Ma et al., 2024). This consistency between approaches lends further support to a (causal) relationship between short-term exposure to NO<sub>2</sub> and PM<sub>2.5</sub> and mortality from specific causes.

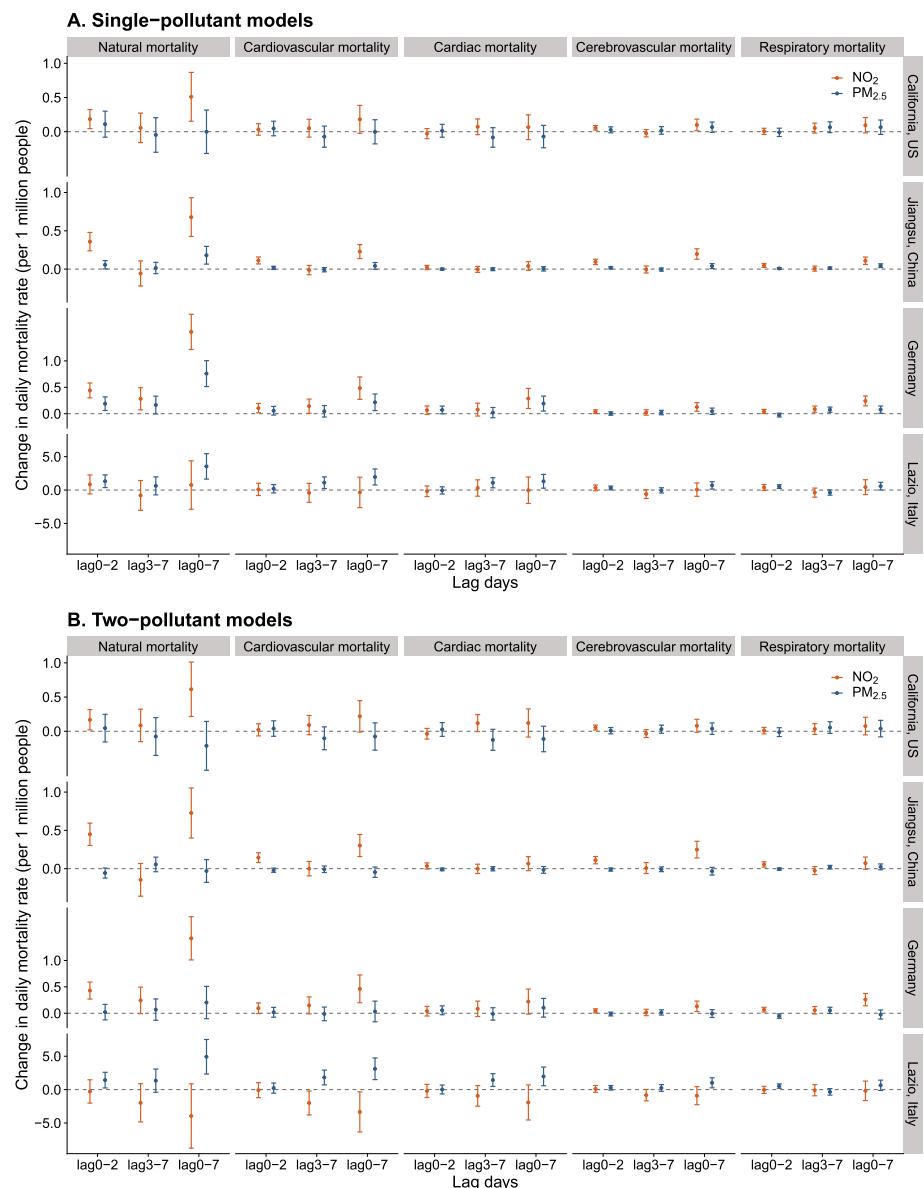
Consistent with previous studies on specific cardiovascular outcomes, we found positive associations with cardiac and cerebrovascular mortality for NO<sub>2</sub> and PM<sub>2.5</sub>. For example, a Chinese study found that each 10 µg/m<sup>3</sup> increase in exposure to NO<sub>2</sub> and PM<sub>2.5</sub> was significantly associated with a 1.46 % and 4.14 % increase in odds of myocardial infarction mortality (Liu et al., 2021). In two-pollutant models, mutually adjusting for NO<sub>2</sub> and PM<sub>2.5</sub>, a Swiss study found a strong association between the eight-day average (lag0-7) of NO<sub>2</sub> and stroke mortality (odds ratio (OR): 1.13, per 10 µg/m<sup>3</sup> increase in NO<sub>2</sub>), particularly pronounced in ischemic (OR: 1.55) compared to hemorrhagic (OR: 1.03) stroke outcomes (Saucy et al., 2021). No association was found for PM<sub>2.5</sub>. Chiusolo et al. reported positive associations between short-term NO<sub>2</sub> and cardiac and stroke mortality at lag0-5 for ten Italian cities (Chiusolo et al., 2011). However, a nationwide Italian study could not confirm these associations for NO<sub>2</sub> but identified a positive association between PM<sub>2.5</sub> and cardiac mortality (Gariazzo et al., 2023). While the association of PM<sub>2.5</sub> with cardiovascular mortality has been classified as causal by the WHO (WHO, 2021), the current evidence on the causal relationship between short-term NO<sub>2</sub> exposure and cardiovascular effects is limited due to insufficient studies (Orellano et al., 2020) and the remaining uncertainty about the independent effects of NO<sub>2</sub> on health (Mills et al., 2016; Forastiere and Peters, 2021; Saucy et al., 2021).

There has been a long discussion on whether NO<sub>2</sub> is a surrogate of other air pollutants, particularly PM<sub>2.5</sub>, that can be co-emitted from the

same sources (Brunekreef et al., 2021; Forastiere and Peters, 2021), and whether exposure to NO<sub>2</sub> independently causes adverse health effects (Mills et al., 2016; Forastiere and Peters, 2021; Saucy et al., 2021). Although statistically adjusting the NO<sub>2</sub> association for PM<sub>2.5</sub> cannot conclusively show an independent effect, gathering evidence might enhance our comprehension of the relationship (Stafoggia et al., 2017). In our study, associations between NO<sub>2</sub> and cause-specific mortality outcomes were generally robust to adjustment for PM<sub>2.5</sub> and, therefore, add to the supporting evidence for independent effects of short-term exposure to NO<sub>2</sub> on mortality (Mills et al., 2016).

To date, strong evidence exists regarding the respiratory effects of short-term exposure to NO<sub>2</sub>, e.g., indicating that inhaling NO<sub>2</sub> can induce allergic inflammation, airway responsiveness, and oxidative stress, thereby triggering asthma attacks (U.S. EPA, 2016). However, the biological pathways underlying the extrapulmonary effects of NO<sub>2</sub> are not yet well understood, and the evidence does not distinctly describe independent NO<sub>2</sub> effects on biological processes leading to mortality (U. S. EPA, 2016; Forastiere and Peters, 2021). For cardiovascular effects, findings from experimental studies indicate changes in heart rate variability, increases in markers of inflammation and oxidative stress, and thrombin generation in plasma of humans and heart tissue of rats as potential nonspecific effects leading to myocardial infarction (Huang et al., 2012; Strak et al., 2013; U.S. EPA, 2016).

Three main pathways are hypothesized for how PM can exert effects on the cardiovascular system (Brook et al., 2010; Rückerl et al., 2011): 1) particles in the lung can induce subclinical systemic reactions such as the release of pro-inflammatory and pro-oxidative mediators leading to several local and systemic inflammatory processes, promoting endothelial dysfunction, a pro-coagulation state, and triggering atherosclerotic plaque; 2) particles settling within the pulmonary tree can directly stimulate neuronal reflexes, causing alterations in pulmonary and cardiac autonomic regulation. These changes in autonomic tone are often the most immediate reaction to the inhalation of air pollution and involve multiple reflex arcs; 3) after exposure, PM (i.e., ultrafine



**Fig. 2.** Estimated change in daily cause-specific mortality rate (per 1 million people) associated with a  $10 \mu\text{g}/\text{m}^3$  increase in  $\text{NO}_2$  or  $\text{PM}_{2.5}$  concentration. This figure shows the estimated change in daily cause-specific mortality rate (per 1 million people) per  $10 \mu\text{g}/\text{m}^3$  increase in  $\text{NO}_2$  or  $\text{PM}_{2.5}$  concentration from single-pollutant models (A) and two-pollutant models (B) on different lag days in Jiangsu, China; California, US; Germany; and Lazio, Italy (2015–2019). The error bars represent the 95 % confidence intervals. Due to the large differences in the magnitude of the confidence intervals, the y-axis for Lazio shows a different range to ensure readability.

Abbreviations as in Fig. 1.

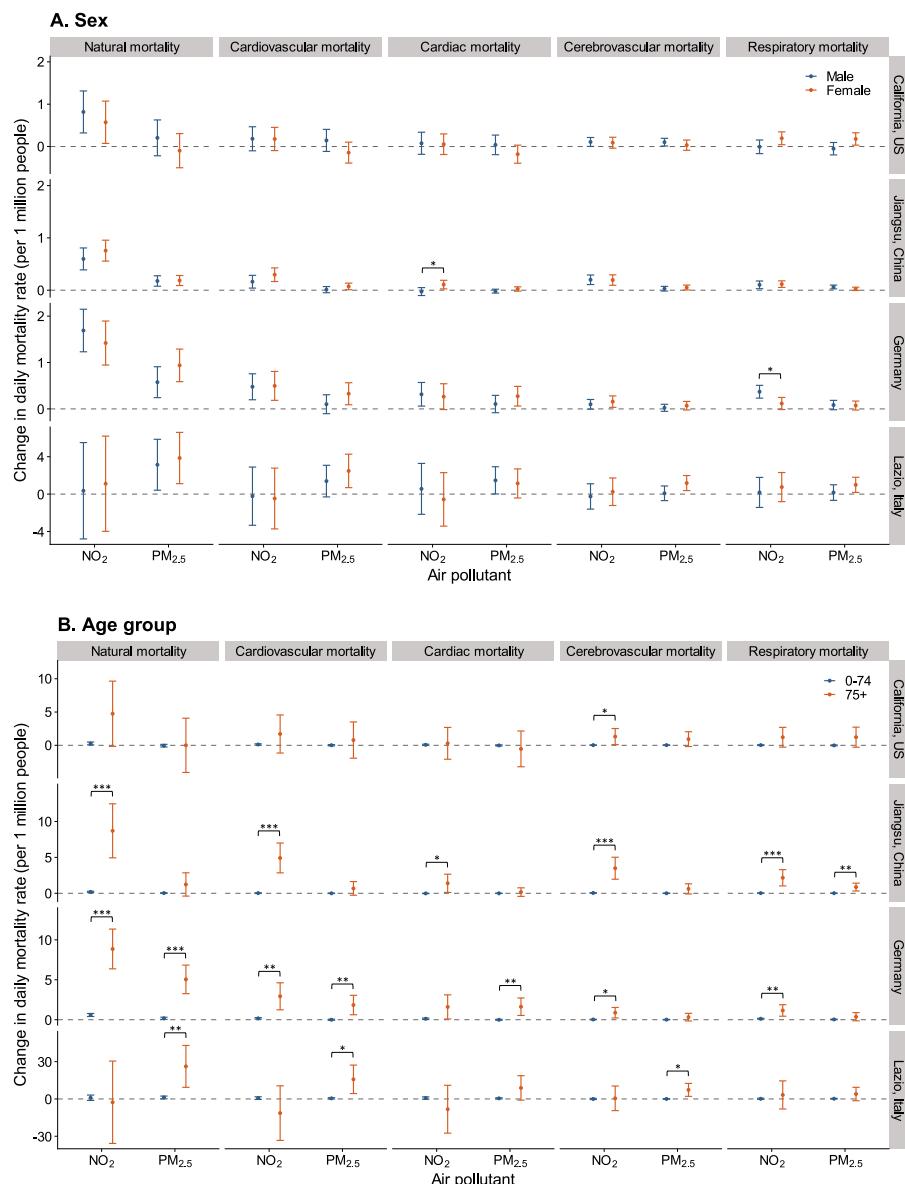
particles) or particle constituents may quickly pass from the pulmonary epithelium into the circulation and interact directly with the cardiovascular system. These small particles may impact the vascular endothelium and atherosclerotic plaques and contribute to oxidative stress and local inflammation.

Our study highlighted elevated effect estimates at prolonged lags (lag0-7) observed across all studied cause-specific outcomes and study regions. Comparable lag patterns were observed in prior studies (Chiusolo et al., 2011; Saucy et al., 2021; Gariazzo et al., 2023), suggesting that increases in air pollution over several days might have particularly adverse impacts on mortality.

Higher risk estimates among older people have been reported previously (Liu et al., 2022; Gariazzo et al., 2023), which we confirmed in our study. These results reflect both the higher mortality rate and the increased susceptibility to short-term air pollution exposure among older people, probably due to reduced compensatory processes and a

higher prevalence of comorbidities in the older population (Shumake et al., 2013).

Previous studies on the acute effects of air pollution on cause-specific mortality primarily focused on metropolitan areas (Linares et al., 2018; Liu et al., 2019; Meng et al., 2021). Our analyses encompassed both urban and rural areas, which is crucial as a substantial share of the population in our study regions resided in non-urban environments which may differ in their emission sources and pollutant concentrations. Recent studies conducted in Sweden (Stafoggia et al., 2020), Italy (Renzi et al., 2022), and the US (Kloog et al., 2014; Bravo et al., 2017) have provided evidence of harmful short-term effects of air pollution in rural and less urbanized areas. Corroborating prior findings, we showed that effect estimates from urban areas barely differed from those of rural areas, even though  $\text{NO}_2$  concentrations in urban areas were considerably higher than in rural ones. Lacking data on particle composition, we can only speculate that the composition and potentially differential toxicity



**Fig. 3.** Estimated change in daily cause-specific mortality rate (per 1 million people) associated with a  $10 \mu\text{g}/\text{m}^3$  increase in  $\text{NO}_2$  or  $\text{PM}_{2.5}$  concentration by effect modification.

This figure shows the estimated change in daily cause-specific mortality rate (per 1 million people) per  $10 \mu\text{g}/\text{m}^3$  increase in  $\text{NO}_2$  or  $\text{PM}_{2.5}$  concentration by sex (male, female) (A), by age group (0–74,  $\geq 75$  years) (B), and urbanicity (urban, rural areas) (C) from single-pollutant models on lag0-7 in Jiangsu, China; California, US; Germany; and Lazio, Italy (2015–2019). The error bars represent the 95 % confidence intervals. Due to the large differences in the magnitude of the confidence intervals, the y-axis for Lazio shows a different range to ensure readability. We tested the statistical differences in effect estimates between groups based on the z score calculated using the coefficients and standard errors for different groups. Significant differences are marked as follows:

\* $p < 0.05$ ; \*\* $p < 0.01$ ; \*\*\* $p < 0.001$ .

Abbreviations as in Fig. 1.

of the particles (Kloog et al., 2014) have a different impact on the short-term association between air pollution and cause-specific mortality in urban and rural settings.

In Jiangsu, a region with high air pollution concentrations, we found that changes in daily  $\text{PM}_{2.5}$  were accompanied by weaker associations with daily changes in mortality rate compared to Germany and Lazio, which have lower levels of air pollution. This finding has been reported in previous studies (Peters et al., 2000; Chen et al., 2012; Chen et al., 2017; Liu et al., 2019). It is hypothesized that in regions with higher long-term exposure to  $\text{PM}_{2.5}$ , populations may exhibit adaptive responses that could result in smaller estimate-per-unit changes in exposure (Peters et al., 2000; Liu et al., 2019).

This comprehensive multi-country study had several strengths, including applying standardized approaches to data collection and analyses in multiple locations. The availability of national, state, or regional cause-specific mortality data matched with highly resolved spatiotemporal exposure models (monitoring sites in California) allowed us to study large portions of the population in the four regions, providing estimates of associations for both urban and rural areas, the latter often neglected in former epidemiological studies. The large sample size and resulting statistical power enabled us to investigate cardiac and cerebrovascular deaths. Applying the novel IFE model allowed us to provide robust support for a causal association between short-term exposure to  $\text{NO}_2$  and  $\text{PM}_{2.5}$  and cause-specific mortality.

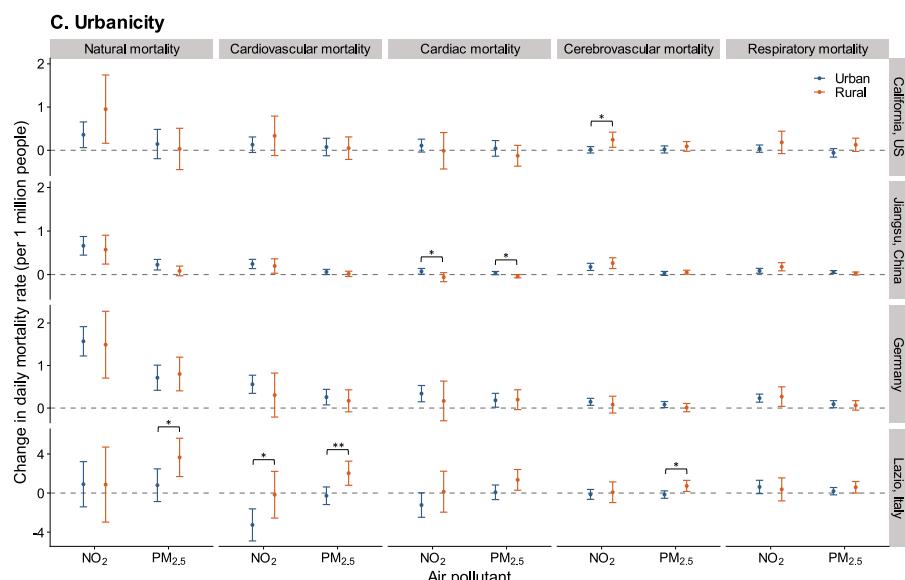


Fig. 3. (continued).

We acknowledge several limitations of this study. First, performing multiple analyses, we cannot rule out that some of our results are observed due to chance. Second, using air pollution data aggregated on a spatial unit level, this study could not account for variations within individual units. Despite high-quality air pollution models, the exposure estimates were subject to uncertainty. In California, only including counties with air quality monitoring stations did not allow us to attain full spatial coverage. Moreover, in California counties with multiple monitoring sites, average air pollution concentrations were calculated at the county level. However, no further spatial interpolation was conducted. Third, we acknowledge the occurrence of misclassification of cause of death statements on death certificates, which could bias our results. Fourth, due to data protection regulations, we were only able to analyze two age groups (0–74 years and  $\geq 75$  years), even though it might have been informative to examine more granular age groups in the 0–74 years range, especially children and adolescents, who are still developing and may be particularly vulnerable to the effects of poor air quality. Last, including four regions from three continents, we observed spatial heterogeneity in the association between short-term changes in air pollution and daily mortality rates across regions, which might be due to different PM components, long-term air pollution levels, regional climate, and population characteristics affecting susceptibility.

## 5. Conclusion

This study, using a causal modeling approach and including urban and rural populations, found that increased short-term  $\text{NO}_2$  and  $\text{PM}_{2.5}$  exposures were associated with increased mortality rates due to natural, cardiovascular, cardiac, cerebrovascular, and respiratory causes. The associations were particularly pronounced among individuals aged 75 and older and with prolonged exposures and remained consistent even in two-pollutant models. These findings emphasize the need of further improving ambient air pollution levels to yield greater public health benefits.

## CRediT authorship contribution statement

**Anne Marb:** Writing – original draft, Visualization, Validation, Methodology, Investigation, Formal analysis, Data curation. **Yiqun Ma:** Writing – review & editing, Visualization, Validation, Investigation, Formal analysis, Data curation. **Federica Nobile:** Writing – review & editing, Visualization, Validation, Investigation, Formal analysis, Data

curation. **Robert Dubrow:** Writing – review & editing, Methodology, Investigation. **Patrick L. Kinney:** Writing – review & editing, Methodology, Investigation. **Massimo Stafoggia:** Writing – review & editing, Supervision, Project administration, Methodology, Investigation, Data curation, Conceptualization. **Kai Chen:** Writing – review & editing, Supervision, Project administration, Methodology, Investigation, Funding acquisition, Data curation, Conceptualization. **Annette Peters:** Writing – review & editing, Supervision, Methodology. **Susanne Breitner:** Writing – review & editing, Validation, Supervision, Project administration, Methodology, Investigation, Data curation, Conceptualization.

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## Declaration of competing interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests:

Anne Marb reports financial support was provided by Health Effects Institute and by Deutsche Forschungsgemeinschaft. Yiqun Ma, Federica Nobile, Robert Dubrow, Kai Chen, and Susanne Breitner report financial support support was provided by Health Effects Institute. If there are other authors, they declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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## Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.envpol.2025.126059>.

## Data availability statement

The modeled air pollution data in Jiangsu (China), Germany, and Lazio (Italy) are available upon request, the station-based air pollution data in California (US) are publicly available on the US EPA website (<https://www.epa.gov/outdoor-air-quality-data>). The cause-specific mortality data in all four regions are confidential.

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