



LUDWIG-  
MAXIMILIANS-  
UNIVERSITÄT  
MÜNCHEN

INSTITUT FÜR STATISTIK  
SONDERFORSCHUNGSBEREICH 386



Czado, Haug:

## A fractionally integrated ECOGARCH process

Sonderforschungsbereich 386, Paper 484 (2006)

Online unter: <http://epub.ub.uni-muenchen.de/>

Projektpartner



## A FRACTIONALLY INTEGRATED ECOGARCH PROCESS

STEPHAN HAUG<sup>\* \*\*</sup> AND  
CLAUDIA CZADO,<sup>\*\*</sup> *Munich University of Technology*

### Abstract

In this paper we introduce a fractionally integrated exponential continuous time GARCH( $p, d, q$ ) process. It is defined in such a way that it is a continuous time extension of the discrete time FIEGARCH( $p, d, q$ ) process. We investigate stationarity and moment properties of the new model. It is also shown that the long memory effect introduced in the log-volatility propagates to the volatility process.

*Keywords:* fractionally integrated exponential continuous time GARCH process; long memory FIEGARCH; ECOGARCH; Lévy process; stationarity; stochastic volatility

2000 Mathematics Subject Classification: Primary 60G10, 60G12, 91B70  
Secondary 91B84

### 1. Introduction

GARCH type processes have become very popular in financial econometrics to model returns of stocks, exchange rates and other series observed at equidistant time points. They have been designed (see Engle [12] and Bollerslev [6]) to capture so-called *stylized facts* of such data, which are e.g. volatility clustering, dependence without correlation and tail heaviness. Several authors have found empirical evidence for the existence of long-run volatility persistence in financial data. Among these are e.g. Andersen and Bollerslev [1], who analysed high-frequency foreign exchange data, Baillie et al. [5], Bollerslev and Mikkelsen [7] and Baillie [4], who gave an overview over long memory processes in econometrics. These findings have to be treated carefully since certain empirical evidence, like a slow decay of the empirical autocorrelation function, could also be due to non-stationarity of the data. This was e.g. shown by Mikosch and Stărică [19] for a long time series of S&P 500 log-returns. In the following this problem will not be our subject.

Since there are different ways to characterise long range dependence, we first want to recall the definition of a long memory process as we will use it before we go on.

#### Definition 1.1.

Let  $Z$  be a stationary stochastic process and  $\gamma_Z(h) = \text{Cov}(Z_{t+h}, Z_t)$ ,  $h \in \mathbb{R}$ , be its

---

\* Postal address: Center for Mathematical Sciences, Munich University of Technology, D-85747 Garching, Germany

\* Email address: haug@ma.tum.de

\*\* Email address: cczado@ma.tum.de

autocovariance function. If there exists  $0 < d < 0.5$  and a constant  $c_Z > 0$  such that

$$\lim_{h \rightarrow \infty} \frac{\gamma_Z(h)}{h^{2d-1}} = c_Z, \quad (1.1)$$

then  $Z$  is called a stationary process with long memory.

In the discrete time GARCH framework there are various models with long range dependence in the volatility process. Among these are the IGARCH( $p, q$ ) process of Engle and Bollerslev [13], the FIGARCH( $p, d, q$ ) process proposed by Baillie et al. [5] or the fractionally integrated EGARCH( $p, d, q$ ) process of Bollerslev and Mikkelsen [7]. The FIGARCH process has to be treated carefully since the existence of a stationary version is not clear; see section 4 in Mikosch and Stărică [19] and Remark 3.2 in Kazakevičius and Leipus [15]. The FIEGARCH( $p, d, q$ ) process is a modification of the EGARCH model of Nelson [21] in the sense that the log-volatility process is modeled by a fractionally integrated ARMA( $p, d, q$ ) process instead of a short memory ARMA process. This long memory effect introduced in the log-volatility process propagates to the volatility and the squared return process. This was shown by Surgailis and Viano [25].

The availability of high frequency data, which increased enormously in the last years, is one reason to consider continuous time models with similar behaviour as discrete time GARCH models. The reason for this is ofcourse that at the highest available frequency the observations of the price process occur at irregularly spaced time points and therefore it is kind of natural to assume an underlying continuous time model. Different approaches have been taken to set up a continuous time model, which has the same features as discrete time GARCH processes. Recently Klüppelberg et al. [16] developed a continuous time GARCH(1, 1) model, which was extended by Brockwell et al. [9] to a continuous time GARCH( $p, q$ ) process for general orders  $p, q \in \mathbb{N}$ ,  $q \geq p$ , henceforth called COGARCH( $p, q$ ). Their approach differs fundamentally from previous attempts, which could be summarized as diffusion approximations (see e.g. Nelson [20]), by the fact that their model is driven by only one source of randomness (like discrete time GARCH) instead of two (like in the diffusion approximations). Haug and Czado [14] have defined an exponential continuous time GARCH process, which is a continuous time extension to the EGARCH process. All these models exhibit a short memory in the volatility process. To incorporate a long memory effect into a continuous time model Comte and Renault [11] defined a continuous time stochastic volatility (SV) model by specifying the log-volatility process as an OU process driven by a fractional Brownian motion. Brockwell and Marquardt [10] proposed to model the stochastic volatility as a non-negative fractionally integrated CARMA process. Another non-Gaussian continuous time SV model with long memory was introduced by Anh et al. [2], where they define their model via the Green function solution of a fractional differential equation driven by a Lévy process. Since this shows considerable interest in continuous time models with long memory in the volatility process, we now want to show in this paper how to extend the ECOGARCH( $p, q$ ) in such a way.

The paper is now organized as follows. In section 2 we define the *fractionally integrated exponential continuous time* GARCH( $p, d, q$ ) process after a short review of elementary properties of Lévy processes and analyse stationarity conditions. The second order behaviour of the volatility process is investigated in section 3, while section 4 deals with second order behaviour of the return process.

## 2. Fractionally integrated exponential COGARCH

In this section we want to construct a continuous time analogue of the discrete time fractionally integrated EGARCH( $p, d, q$ ) process, which is defined in the following way:

Let  $p, q \in \mathbb{N}, \mu, \alpha_1, \dots, \alpha_q, \beta_1, \dots, \beta_p \in \mathbb{R}$ , suppose  $\alpha_q \neq 0, \beta_p \neq 0$  and that the autoregressive polynomial

$$\phi(z) := 1 - \alpha_1 z - \dots - \alpha_q z^q$$

and the moving average polynomial

$$\psi(z) := \beta_1 + \beta_2 z + \dots + \beta_p z^{p-1}$$

have no common zeros and that  $\phi(z) \neq 0$  on  $\{z \in \mathbb{C} \mid |z| \leq 1\}$ . Let  $(\epsilon_n)_{n \in \mathbb{Z}}$  be an i.i.d. sequence with  $\mathbb{E}(\epsilon_1) = 0$  and  $\text{Var}(\epsilon_1) = 1$  and  $-0.5 < d < 0.5$ . Define the measurable function  $f : \mathbb{R} \rightarrow \mathbb{R}$  by

$$f(x) := \theta x + \gamma[|x| - \mathbb{E}(|x|)], \quad x \in \mathbb{R}, \quad (2.1)$$

with real coefficients  $\theta$  and  $\gamma$ . Then we call  $(X_n)_{n \in \mathbb{Z}}$ , where  $X_n = \sigma_n \epsilon_n$ , a FIE-GARCH( $p, d, q$ ) process if

$$\log(\sigma_n^2) = \mu + \phi(B)^{-1}(1 - B)^{-d}(1 + \psi(B))f(\epsilon_{n-1}), \quad (2.2)$$

where  $B$  is the backward shift operator,  $BX_n = X_{n-1}$ .

We will define the process using the idea of Klüppelberg et al. [16] to replace the innovations  $\epsilon_n$  of the discrete time model by the jumps of a Lévy process  $L = (L_t)_{t \geq 0}$ . Any Lévy process  $L$  on  $\mathbb{R}$  has a characteristic function of the form

$$\mathbb{E}(e^{iuL_t}) = \exp\{t\psi_L(u)\}, \quad t \geq 0,$$

with

$$\psi_L(u) := i\gamma_L u - \frac{\tau_L^2}{2} u^2 + \int_{\mathbb{R}} (e^{iux} - 1 - iux\chi_{\{|x| \leq 1\}}) \nu_L(dx), \quad u \in \mathbb{R},$$

where  $\tau_L^2 \geq 0$ ,  $\gamma_L \in \mathbb{R}$ , the measure  $\nu_L$  satisfies

$$\nu_L(\{0\}) = 0 \quad \text{and} \quad \int_{\mathbb{R}} \min(x^2, 1) \nu_L(dx) < \infty$$

and  $\chi_A(\cdot)$  denotes the indicator function of the set  $A \subset \mathbb{R}$ . The measure  $\nu_L$  is called the Lévy measure of  $L$  and the triplet  $(\gamma_L, \tau_L^2, \nu_L)$  is called the characteristic triplet of  $L$ . The map  $\psi_L$  is called the Lévy symbol or Lévy exponent. For  $L$  to have finite mean and variance it is necessary and sufficient that

$$\int_{|x| > 1} |x| \nu_L(dx) < \infty \quad \text{and} \quad \int_{|x| > 1} x^2 \nu_L(dx) < \infty,$$

respectively (Sato [23], Example 25.12). For more details on Lévy processes we refer to Sato [23] or Applebaum [3].

We consider zero mean Lévy processes  $L$  defined on a probability space  $(\Omega, \mathcal{F}, P)$  with jumps  $\Delta L_t := L_t - L_{t-}$ . Since  $\mathbb{E}(L_t) = t(\gamma_L + \int_{|x|>1} x\nu_L(dx))$ , a zero mean implies that  $\gamma_L = -\int_{|x|>1} x\nu_L(dx)$  and hence the corresponding Lévy symbol is of the form

$$\psi_L(u) = -\tau_L^2 \frac{u^2}{2} + \int_{\mathbb{R}} (e^{iux} - 1 - iux)\nu_L(dx), \quad (2.3)$$

and the Lévy-Itô decomposition (see e.g. Theorem 2.4.16 of Applebaum [3]) of  $L$  is

$$L_t = B_t + \int_{\mathbb{R}-\{0\}} x \tilde{N}_L(t, dx), \quad t \geq 0, \quad (2.4)$$

where  $B$  is a Brownian motion with variance  $\tau_L^2$  and  $\tilde{N}_L(t, dx) = N_L(t, dx) - t\nu_L(dx)$ ,  $t \geq 0$ , is the compensated random measure associated to the Poisson random measure

$$N_L(t, A) = \#\{0 \leq s < t; \Delta L_s \in A\} = \sum_{0 < s \leq t} \chi_A(\Delta L_s), \quad A \in \mathcal{B}(\mathbb{R} - \{0\}),$$

on  $\mathbb{R}_+ \times \mathbb{R} - \{0\}$ , which is independent of  $B$ .

The Lévy process  $L$  can be extended to a Lévy process  $L^*$  defined on the whole real line by choosing a second Lévy process  $(\tilde{L}_t)_{t \geq 0}$  independent of  $L$  and with the same distribution as  $L$  and specifying

$$L_t^* := L_t \chi_{[0, \infty)}(t) - \tilde{L}_{-t} \chi_{(-\infty, 0)}(t), \quad t \in \mathbb{R},$$

where  $\chi_A(\cdot)$  denotes the indicator function of the set  $A$ . In the following we will work with  $L^*$  but write for simplicity  $L$  instead of  $L^*$ .

Now we are able to define the *fractionally integrated exponential continuous time GARCH* ( $p, d, q$ ) process, shortly called FIECOGARCH( $p, d, q$ ). We will see that if the log-volatility process defined it is actually a fractionally integrated continuous time ARMA( $q, d, p-1$ ) process, henceforth called FICARMA( $q, d, p-1$ ) process (see e.g. Brockwell and Marquardt [10] for details on FICARMA processes). The driving noise process of the log-volatility process will be defined similarly to (2.1).

**Definition 2.1.** Let  $L = (L_t)_{t \geq 0}$  be a Lévy process with  $\mathbb{E}(L_1) = 0$ ,  $\mathbb{V}\text{ar}(L_1) = 1$  and Lévy measure  $\nu_L$  and let the  $q \times q$  matrix  $A$  and vectors  $\mathbf{b} \in \mathbb{R}^q$  and  $\mathbf{1}_q \in \mathbb{R}^q$  be defined by

$$A = \begin{bmatrix} 0 & 1 & 0 & \cdots & 0 \\ 0 & 0 & 1 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & 1 \\ -a_q & -a_{q-1} & -a_{q-2} & \cdots & -a_1 \end{bmatrix}, \quad \mathbf{b} = \begin{bmatrix} b_1 \\ b_2 \\ \vdots \\ b_{q-1} \\ b_q \end{bmatrix}, \quad \mathbf{1}_q = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \\ 1 \end{bmatrix}$$

with coefficients  $a_1, \dots, a_q, b_1, \dots, b_p \in \mathbb{R}$ , where  $a_q \neq 0, b_p \neq 0$ , and  $b_{p+1} = \dots = b_q = 0$ . Then for  $0 < d < 0.5$  we define the fractionally integrated exponential COGARCH( $p, d, q$ ) process  $G_d$  as the stochastic process satisfying,

$$dG_{d,t} = \sigma_{d,t} dL_t, \quad t > 0, \quad G_0 = 0, \quad (2.5)$$

where the log-volatility process is given by

$$\log(\sigma_{d,t+}^2) = \mu + \int_{-\infty}^t g_d(t-u) dM_u, \quad t > 0, \quad (2.6)$$

with mean  $\mu \in \mathbb{R}$  and initial value  $\log(\sigma_{d,0}^2)$  independent of the driving Lévy process  $L$ . The process

$$M_t := \int_{\mathbb{R}-\{0\}} h(x) \tilde{N}_L(t, dx), \quad t > 0, \quad (2.7)$$

is a zero mean Lévy process (see Remark 2.2) with

$$h(x) := \theta x + \gamma |x|$$

and parameters  $\theta, \gamma \in \mathbb{R}$ . The kernel function

$$g_d(t) = \int_0^t g(t-u) \frac{u^{d-1}}{\Gamma(d)} du, \quad 0 < d < 0.5, \quad (2.8)$$

is the Riemann-Liouville fractional integral of order  $d$  (see Definition 2.1 in Sako et al. [22]) of the kernel function  $g(t) = \mathbf{b}^T e^{At} \mathbf{1}_q \chi_{(0,\infty)}(t)$ .

Returns over a time interval of length  $r > 0$  are described by the increments of  $G$

$$G_{d,t}^{(r)} := G_{d,t} - G_{d,t-r} = \int_{(t-r,t]} \sigma_{d,s} dL_s, \quad t \geq r > 0. \quad (2.9)$$

On the other hand an equidistant sequence of such non-overlapping returns of length  $r$  is given by  $(G_{nr}^{(r)})_{n \in \mathbb{N}}$ . Thus this gives us the possibility to model ultra high frequency data, which consists of returns over varying time intervals.

In the rest of the paper the following terminology will be used:

$G_d$	(log-) price process
$G_d^{(r)}$	(log-) return process
$\sigma_d^2$	volatility process
$\log(\sigma_d^2)$	log-volatility process.

**Remark 2.2.** (i) The log-volatility process (2.6) is well-defined and stationary if the real part of the eigenvalues of  $A$  is negative, since then

$$\int_{\mathbb{R}} \int_{\mathbb{R}-\{0\}} |g_d(t-s)x|^2 \nu_L(dx) ds, \quad \forall t \geq 0,$$

and we can apply Theorem 4.3.4 and 4.3.16 in Applebaum [3] from which the assumptions follow. In this case the log-volatility process is a FICARMA( $p, d, q$ ) process.

(ii) The process  $M$  defined by (2.7) is by construction a process with independent and stationary increments and by Theorem 4.3.4 in Applebaum [3] well defined if

$$\int_{\mathbb{E}} |h(x)|^2 \nu_L(dx) < \infty. \quad (2.10)$$

Condition (2.10) is satisfied since  $\nu_L$  is a Lévy measure and  $L$  has finite variance. By equation (2.9) of Applebaum [3] the characteristic function of  $M$  at time  $t \geq 0$  is given by

$$\begin{aligned} \mathbb{E}(e^{iuM_t}) &= \exp\left(t \int_{\mathbb{R}} [e^{iux} - 1 - iux] \nu_M(dx)\right) \\ &= \exp\left(t \left\{ iu\gamma_M + \int_{\mathbb{R}} [e^{iux} - 1 - iux\chi_{|x|\leq 1}] \nu_M(dx) \right\}\right) \\ &=: \exp(t\psi_M(u)), \end{aligned}$$

where  $\nu_M := \nu_L \circ h^{-1}$  is the Lévy measure of  $M$  and  $\gamma_M := -\int_{|x|>1} x\nu_M(dx)$ . The concrete form of  $\nu_M$  depends on the sign and size of  $\theta$  and  $\gamma$  and is given in the following:

$$\nu_M((-\infty, -x]) = \begin{cases} \nu_L([\frac{-x}{\theta+\gamma}, \infty)) + \nu_L((-\infty, -\frac{x}{\theta-\gamma}]), & \theta + \gamma < 0 \text{ and } \theta - \gamma > 0 \\ \nu_L((-\infty, -\frac{x}{\theta-\gamma}]), & \theta - \gamma > 0 \text{ and } \theta + \gamma > 0 \\ \nu_L([\frac{-x}{\theta+\gamma}, \infty)), & \theta + \gamma < 0 \text{ and } \theta - \gamma < 0 \\ 0 & \theta + \gamma > 0 \text{ and } \theta - \gamma < 0 \end{cases}$$

and

$$\nu_M([x, \infty)) = \begin{cases} \nu_L([\frac{x}{\theta+\gamma}, \infty)) + \nu_L((-\infty, \frac{x}{\theta-\gamma}]), & \theta + \gamma > 0 \text{ and } \theta - \gamma < 0 \\ \nu_L((-\infty, \frac{x}{\theta-\gamma}]), & \theta - \gamma < 0 \text{ and } \theta + \gamma < 0 \\ \nu_L([\frac{x}{\theta+\gamma}, \infty)), & \theta + \gamma > 0 \text{ and } \theta - \gamma > 0 \\ 0 & \theta + \gamma < 0 \text{ and } \theta - \gamma > 0 \end{cases}$$

for  $x > 0$ . One recognises that for  $\theta + \gamma < 0 \vee \theta - \gamma > 0$   $M$  is a spectrally negative Lévy process, i.e.  $M$  has only negative jumps, and for  $\theta + \gamma > 0 \vee \theta - \gamma < 0$   $M$  is a spectrally positive Lévy process. Therefore  $M$  has the characteristic triplet  $(\gamma_M, 0, \nu_M)$ .

(iii) The model can of course also be defined for a different choice of  $h$ , as long as condition (2.10) is satisfied.

Alternatively the log-volatility process can be defined in terms of the fractional Lévy process  $M_d$  associated with  $M$ . We recall the definition of a fractional Lévy process from Marquardt [18].

**Remark 2.3.** Let  $M = (M_t)_{t \in \mathbb{R}}$  be a Lévy process on  $\mathbb{R}$  with  $\mathbb{E}M_1 = 0, \mathbb{E}M_1^2 < \infty$  and without Brownian component. For the fractional integration parameter  $0 < d < 0.5$  the stochastic process

$$M_{d,t} = \frac{1}{\Gamma(d+1)} \int_{\mathbb{R}} [(t-s)_+^d - (-s)_+^d] dM_s, \quad t \in \mathbb{R}, \quad (2.11)$$

is called a fractional Lévy process.

The strictly stationary log-volatility process (2.6) is then equal to

$$\mu + \int_{-\infty}^t \mathcal{D}_+^d g_d(t-u) dM_{d,u}, \quad t > 0, \quad (2.12)$$

in the  $L^2$ -sense, where  $\mathcal{D}_+^d g_d(x) = \frac{1}{\Gamma(1-d)} \frac{d}{dx} \int_{-\infty}^x \frac{g_d(u)}{(x-u)^d} du$  is the Riemann-Liouville fractional derivative of  $g_d$  of order  $d$  (see Definition 2.2 in Sako et al. [22]). Since  $g \in L^1(\mathbb{R})$  we get from Theorem 2.4 in Sako et al. [22] that  $\mathcal{D}_+^d g_d = g$ . The proof of the equivalence of (2.6) and (2.12) can be found in Marquardt [18], Theorem 6.5.

If the Lévy process  $M$  is of finite activity, i.e.  $\nu_M(\mathbb{R}) < \infty$ , then the corresponding fractional Lévy process  $M_d$  is of finite variation. In this case the integral in (2.12) can be defined as a Riemann-Stieltjes integral. In case where  $M$  is not of finite activity the corresponding fractional Lévy process is not a semimartingale, but for a deterministic integrand the integral with respect to  $M_d$  can be defined in the  $L^2$ -sense (we refer to section 5 of Marquardt [18] for details). We do not restrict the driving Lévy process to be of finite activity but we only deal with deterministic integrands and hence this turns out to be sufficient for our purpose.

The log-volatility process (2.12) is now the solution of the continuous time state space model

$$\log(\sigma_{d,t}^2) = \mu + \mathbf{b}^T \mathbf{X}_{d,t-}, \quad t > 0, \quad \log(\sigma_{d,0}^2) = \mu + \mathbf{b}^T \mathbf{X}_{d,0} \quad (2.13)$$

$$d\mathbf{X}_{d,t} = A\mathbf{X}_{d,t}dt + \mathbf{1}_q dM_{d,t}, \quad t > 0, \quad (2.14)$$

where  $\mathbf{X}_{d,0}$  is independent of  $(M_{d,t})_{t \geq 0}$  and  $A$ ,  $\mathbf{b}$  and  $\mathbf{1}_q$  are defined in Definition 2.1.

The state space representation of the log-volatility process is also advantageous for the purpose of simulating the log-price process  $G_d$ . The simulation procedure is the following:

1. Choose simulation times  $0 = t_0 < t_1 < \dots < t_n \leq T$ , possibly random.
2. Generate increments  $M_{d,t_{i+1}} - M_{d,t_i}$ ,  $i = 0, \dots, n-1$ , of the driving fractional Lévy process.
3. Approximate the state process (2.14) of the log-volatility by a stochastic Euler scheme.
4. Compute

$$\log(\widehat{\sigma}_{d,t_i}^2) = \mu + \mathbf{b}^T \widehat{\mathbf{X}}_{d,t_{i-1}}$$

for  $i = 1, \dots, n$ .

5. Compute an approximation  $\widehat{G}_d$  by a stochastic Euler scheme:

$$\widehat{G}_{d,t_i} = \widehat{G}_{d,t_{i-1}} + \widehat{\sigma}_{d,t_{i-1}} W_i + \widehat{\sigma}_{d,t_{i-1}} J_i,$$

where  $W_i \sim N(0, t_i - t_{i-1})$  and  $J_i$  is an increment of the jump part of  $L$  over the time interval  $[t_{i-1}, t_i]$ .



Since the fractional Lévy process  $M_d$  at time  $t$  is an integral with respect to the driving Lévy process  $L$  it can be approximated by the corresponding Riemann sums. This approximation is explained in chapter 2.4.3 in Marquardt [17].

Defined in this way  $\log(\sigma_d^2)$  is not strictly stationary by definition. The conditions for stationarity of  $\log(\sigma_d^2)$  the volatility process  $\sigma_d^2$  and the return process  $G_d^{(r)}$  are summarized in the next proposition. The autocovariance function of the log-volatility process and its asymptotic behaviour is also stated. Therefore we will call two functions  $f_1$  and  $f_2$  asymptotically equivalent if  $\lim_{x \rightarrow \infty} f_1(x)/f_2(x) = 1$  and denote it by  $f_1(x) \sim f_2(x)$  as  $x \rightarrow \infty$ .

**Proposition 2.4.** *Let  $\log(\sigma_d^2)$  be defined by (2.13) and (2.14) and  $G_d$  as in Definition 2.1. If the eigenvalues of  $A$  all have negative real parts and  $\log(\sigma_{d,0}^2)$  has the same distribution as  $\int_0^\infty \mathbf{b}^T e^{As} \mathbf{1}_q dM_{d,s}$ , then  $\log(\sigma_d^2)$  and  $\sigma_d^2$  are strictly stationary and  $G_d$  has strictly stationary increments. The log-volatility process is weakly stationary and has autocovariance function*

$$\begin{aligned} \text{Cov}(\log(\sigma_{d,t+h}^2), \log(\sigma_{d,t}^2)) &= \mathbb{E}(M_1^2) \int_{\mathbb{R}} g_d(u+h)g_d(u)du, & t > 0, h \geq 0 \\ &\sim C_1 h^{2d-1}, & \text{as } h \rightarrow \infty, \end{aligned} \quad (2.15)$$

where  $C_1 := \frac{\Gamma(1-2d)}{\Gamma(d)\Gamma(1-d)} \mathbb{E}M_1^2 \left( \int_{\mathbb{R}} g(s)ds \right)^2$ .

The strict stationarity of  $\log(\sigma_d^2)$ ,  $\sigma_d^2$  and the increments of  $G_d$  follows from the same reasoning as in the short memory case (see Proposition 3.3 and Corollary 3.5 in Haug and Czado [14]). The proof of (2.15) and (2.16) is given in Marquardt [18], Theorem 6.7 and 6.6.

**Remark 2.5.** *The asymptotic behaviour of the autocovariance function of a FICARMA process was derived by Brockwell [8]. The result depends on the asymptotic behaviour of the kernel function  $g_d$*

$$g_d(s) \sim \left( \int_{\mathbb{R}} g(x)dx \right) s^{d-1}, \quad \text{for } s \rightarrow \infty, \quad (2.17)$$

which was shown in Brockwell [8], section 4. In the following the constant in (2.17) will be denoted by  $C_2$ .

### 3. Second order properties of the volatility process

Second order properties are now derived under the assumption that the log-volatility process is strictly stationary. The integral in (2.6) is well defined and from (2.5) in Sato and Yamazato [24] it follows that the characteristic function of  $\log(\sigma_{d,t+}^2)$  is given

by

$$\begin{aligned}
\mathbb{E}(e^{iu \log(\sigma_{d,t}^2)}) &= e^{iu\mu} \exp \left\{ \int_0^\infty \psi_M(g_d(s)u) ds \right\} \\
&= e^{iu\mu} \exp \left\{ iu \left[ \int_0^\infty g_d(s) \gamma_M ds + \int_0^\infty \int_{\mathbb{R}} (\chi_{\{|g_d(s)x| \leq 1\}} - \chi_{\{|x| \leq 1\}}) \nu_M(dx) ds \right] \right. \\
&\quad \left. + \int_0^\infty \int_{\mathbb{R}} (e^{iug_d(s)x} - 1 - iug_d(s)x \chi_{\{|g_d(s)x| \leq 1\}}) \nu_M(dx) ds \right\} \\
&= e^{iu\mu} \exp \left\{ iu\gamma_\infty + \int_{\mathbb{R}} (e^{iux} - 1 - iux \chi_{\{|x| \leq 1\}}) \nu_\infty(dx) \right\},
\end{aligned}$$

since  $\log(\sigma_{d,t}^2) \stackrel{d}{=} \mu + \int_0^\infty g_d(s) dM_s$ .

The stationary distribution  $F_d$  of  $\log(\sigma_d^2)$  is therefore infinitely divisible with characteristic triplet  $(\gamma_{d,\infty}, 0, \nu_{d,\infty})$ , where

$$\begin{aligned}
\gamma_{d,\infty} &= \int_0^\infty g_d(s) \gamma_M ds \\
&\quad + \int_0^\infty \int_{\mathbb{R}} g_d(s) x (\chi_{\{|g_d(s)x| \leq 1\}} - \chi_{\{|x| \leq 1\}}) \nu_M(dx) ds \quad (3.1)
\end{aligned}$$

$$\nu_{d,\infty}(B) = \int_0^\infty \int_{\mathbb{R}} \chi_B(g_d(s)x) \nu_M(dx) ds, \quad B \in \mathcal{B}(\mathbb{R}). \quad (3.2)$$

The second order behaviour is now summarised in the following proposition.

**Proposition 3.1.** *Let  $\log(\sigma_d^2)$  be strictly stationary with marginal distribution  $F_d$ , where  $F_d$  is infinitely divisible with characteristic triplet  $(\gamma_{d,\infty}, 0, \nu_{d,\infty})$ . The  $k$ -th moment of  $\sigma_{d,t}^2$  is finite, if*

$$k \in K_{d,\infty} = \{s \in \mathbb{R} : \int_{|x|>1} e^{sx} \nu_{d,\infty}(dx) < \infty\}.$$

In this case

$$\Psi_{d,\infty}(k) = \int_0^\infty \Psi_M(g_d(s)k) ds, \quad (3.3)$$

where  $\Psi_M(u) := \psi_M(-iu)$ ,  $u \in \mathbb{R}$ , is well defined and

$$\mathbb{E}(\sigma_{d,t}^{2k}) = e^{\mu k} e^{\Psi_{d,\infty}(k)}, \quad \forall t \geq 0. \quad (3.4)$$

Assume that  $\mathbb{E}(\sigma_{d,t}^4) < \infty$ . Let  $\Psi_{d,\infty}^h(k)$  and  $\Psi_d^h(k)$  be defined by (3.3) with kernel function  $g_d$  replaced by

$$g_{d,\infty}^h(s) := g_d(s) + g_d(s+h) \quad \text{and} \quad g_d^h(s) := g_d(s) \chi_{(0,h)}(s)$$

respectively. Then the autocovariance function of  $\sigma_d^2$  is given by

$$\text{Cov}(\sigma_{d,t+h}^2, \sigma_{d,t}^2) = e^{2\mu} (e^{\Psi_{d,\infty}^h(1)} e^{\Psi_d^h(1)} - e^{2\Psi_{d,\infty}(1)}). \quad (3.5)$$

If we replace the kernel functions in the proof of Proposition 4.3 and 4.5 in Haug and Czado [14] appropriately with the kernel functions  $g_d, g_{d,\infty}^h$  and  $g_d^h$ , then the result follows.

Next we will show that the long memory property introduced in the log volatility process implies also a long memory effect in the volatility process. The proof is based on a result for the FICARMA( $p, d, q$ ) process which can be found in Lemma 1.23 in Marquardt [17].

**Theorem 3.2.** *Let  $\log(\sigma_d^2)$  be the strictly stationary long memory process (2.6) with long memory parameter  $0 < d < 0.5$  and assume that  $2 \in K_{d,\infty}$ . Then  $\mathbb{E}(\sigma_{d,t}^4) < \infty$ ,  $\forall t \geq 0$ , and*

$$\text{Cov}(\sigma_{d,t+h}^2, \sigma_{d,t}^2) \sim e^{2(\mu + \Psi_{d,\infty}(1))} C_1 h^{2d-1}, \quad \text{as } h \rightarrow \infty, \quad (3.6)$$

where  $C_1 = \frac{\Gamma(1-2d)}{\Gamma(d)\Gamma(1-d)} \mathbb{E}M_1^2 \left( \int_{\mathbb{R}} g(s) ds \right)^2$ .

**Proof:** From equation (3.5) it follows that

$$\begin{aligned} \text{Cov}(\sigma_{d,t+h}^2, \sigma_{d,t}^2) &= e^{2\mu} (e^{\Psi_{d,\infty}^h(1)} e^{\Psi_d^h(1)} - e^{2\Psi_{d,\infty}(1)}) \\ &= e^{2\mu} e^{2\Psi_{d,\infty}(1)} (e^{\Psi_{d,\infty}^h(1) + \Psi_d^h(1) - 2\Psi_{d,\infty}(1)} - 1) \\ &= e^{2\mu} e^{2\Psi_{d,\infty}(1)} (\Psi_{d,\infty}^h(1) + \Psi_d^h(1) - 2\Psi_{d,\infty}(1)) \\ &\quad + O((\Psi_{d,\infty}^h(1) + \Psi_d^h(1) - 2\Psi_{d,\infty}(1))^2). \end{aligned}$$

If we can show that  $\Psi_{d,\infty}^h(1) + \Psi_d^h(1) - 2\Psi_{d,\infty}(1) \sim C_1 h^{2d-1}$ , as  $h \rightarrow \infty$ , the result follows. Consider therefore

$$\begin{aligned} &\Psi_{d,\infty}^h(1) - \Psi_{d,\infty}(1) + \Psi_d^h(1) - \Psi_{d,\infty}(1) \\ &= \int_0^\infty \int_{\mathbb{R}} \left\{ e^{g_{d,\infty}^h(s)x} - 1 - g_{d,\infty}^h(s)x - \left[ e^{g_d(s)x} - 1 - g_d(s)x \right] \right\} \nu_M(dx) ds \\ &\quad + \int_0^\infty \int_{\mathbb{R}} \left\{ e^{g_d^h(s)x} - 1 - g_d^h(s)x - \left[ e^{g_d(s)x} - 1 - g_d(s)x \right] \right\} \nu_M(dx) ds \\ &= \int_0^\infty \int_{\mathbb{R}} \left\{ e^{g_{d,\infty}^h(s)x} - e^{g_d(s)x} + e^{g_d^h(s)x} - e^{g_{d,\infty}^h(s)x} \right\} \nu_M(dx) ds \\ &= \int_0^\infty \int_{\mathbb{R}} e^{g_d(s)x} \left\{ e^{g_d(s+h)x} - 1 \right\} \nu_M(dx) ds - \int_h^\infty \int_{\mathbb{R}} \left\{ 1 - e^{g_d(s)x} \right\} \nu_M(dx) ds. \end{aligned}$$

Series expansion of the exponential function yields

$$\begin{aligned}
& \Psi_{d,\infty}^h(1) - \Psi_{d,\infty}(1) + \Psi_d^h(1) - \Psi_d(1) \\
&= \int_0^\infty \int_{\mathbb{R}} \left\{ \sum_{k=1}^{\infty} \frac{[g_d(s+h)x]^k}{k!} + g_d(s)x \sum_{k=1}^{\infty} \frac{[g_d(s+h)x]^k}{k!} \right. \\
&\quad \left. + \sum_{m=2}^{\infty} \frac{(g_d(s)x)^m}{m!} \sum_{k=1}^{\infty} \frac{(g_d(s+h)x)^k}{k!} - \sum_{k=1}^{\infty} \frac{[g_d(s+h)x]^k}{k!} \right\} \nu_M(dx) ds \\
&= \int_0^\infty \int_{\mathbb{R}} \left[ x g_d(s) \sum_{m=1}^{\infty} \frac{(g_d(s+h)x)^m}{m!} + \sum_{m=2}^{\infty} \frac{(g_d(s)x)^m}{m!} \sum_{k=1}^{\infty} \frac{(g_d(s+h)x)^k}{k!} \right] \nu_M(dx) ds \\
&= \int_0^\infty \int_{\mathbb{R}} \left[ x^2 g_d(s) g_d(s+h) + g_d(s)x \sum_{m=2}^{\infty} \frac{(g_d(s+h)x)^m}{m!} \right. \\
&\quad \left. + g_d(s+h)x \sum_{m=2}^{\infty} \frac{(g_d(s)x)^m}{m!} + \sum_{m=2}^{\infty} \frac{(g_d(s)x)^m}{m!} \sum_{k=2}^{\infty} \frac{(g_d(s+h)x)^k}{k!} \right] \nu_M(dx) ds \\
&= \int_0^\infty \int_{\mathbb{R}} \left[ g_d(s+h)x \sum_{m=1}^{\infty} \frac{(g_d(s)x)^m}{m!} + \sum_{m=1}^{\infty} \frac{(g_d(s)x)^m}{m!} \sum_{k=2}^{\infty} \frac{(g_d(s+h)x)^k}{k!} \right] \nu_M(dx) ds.
\end{aligned}$$

Define  $M_j := \int_{\mathbb{R}} x^j \nu_M(dx)$ ,  $j \in \mathbb{N}$ . Since  $\int_{|x|>1} e^x \nu_{d,\infty}(dx) < \infty$  implies that  $\int_{\mathbb{R}} |x|^k \nu_M(dx) < \infty$ ,  $k \geq 2$ , we get that all moments  $M_j$ ,  $j \geq 2$ , of the Lévy measure  $\nu_M$  are finite. Consider now the integral

$$I_1(h) := \int_0^\infty g_d(s+h)g_d(s) \left[ M_2 + \sum_{m=2}^{\infty} \frac{(g_d(s))^{m-1}}{m!} M_{m+1} \right] ds.$$

We want to show that

$$I_1(h) \sim M_2 \int_0^\infty G_d(s+h)G_d(s)ds =: I_G(h), \quad \text{as } h \rightarrow \infty, \quad (3.7)$$

with  $G_d(s) := C_2 s^{d-1}$ , since  $I_G(h) \sim C_1 h^{2d-1}$ ,  $h \rightarrow \infty$ . We show first, that

$$\begin{aligned}
& g_d(s+h+h^{d/2})g_d(s+h^{d/2}) \left[ M_2 + \sum_{m=2}^{\infty} \frac{(g_d(s+h^{d/2}))^{m-1}}{m!} M_{m+1} \right] \\
&\quad \sim M_2 G_d(s+h+h^{d/2})G_d(s+h^{d/2}),
\end{aligned}$$

if  $h \rightarrow \infty$ . Consider therefore the limit

$$\begin{aligned}
& \lim_{s \geq 0, h \rightarrow \infty} \frac{g_d(s+h+h^{d/2})g_d(s+h^{d/2}) \left[ M_2 + \sum_{m=2}^{\infty} \frac{(g_d(s+h^{d/2}))^{m-1}}{m!} M_{m+1} \right]}{M_2 G_d(s+h+h^{d/2})G_d(s+h^{d/2})} \\
&= 1 + \lim_{s \geq 0, h \rightarrow \infty} M_2^{-1} \sum_{m=2}^{\infty} \frac{(g_d(s+h^{d/2}))^{m-1}}{m!} M_{m+1},
\end{aligned}$$

which is equal to 1 because of (2.17),

$$\begin{aligned} & \left| g_d(s + h^{d/2})^{-1} \sum_{m=2}^{\infty} \frac{(g_d(s + h^{d/2}))^m}{m!} M_{m+1} \right| \\ & \leq M^* |g_d(s + h^{d/2})|^{-1} \left( e^{|g_d(s + h^{d/2})|} - |g_d(s + h^{d/2})| - 1 \right), \end{aligned}$$

with  $M^* := \sup_{j \in \mathbb{N}} |M_j| < \infty$ , and  $\lim_{x \geq 0, x \rightarrow 0} x^{-1}(e^x - x - 1) = 0$ . From Lemma 1.22 in Marquardt [17] it follows that

$$\tilde{I}_1(h) \sim \tilde{I}_G(h), \quad \text{for } h \rightarrow \infty, \quad (3.8)$$

where

$$\tilde{I}_1(h) := \int_{h^{d/2}}^{\infty} g_d(s + h) g_d(s) \left[ M_2 + \sum_{m=2}^{\infty} \frac{(g_d(s))^{m-1}}{m!} M_{m+1} \right] ds$$

and

$$\tilde{I}_G(h) := M_2 \int_{h^{d/2}}^{\infty} G_d(s + h) G_d(s) ds.$$

Now (3.7) follows if we can show

$$\frac{|I_1(h) - I_G(h)|}{|I_G(h)|} \leq \frac{|I_1(h) - \tilde{I}_1(h)|}{|I_G(h)|} + \frac{|\tilde{I}_1(h) - \tilde{I}_G(h)|}{|I_G(h)|} + \frac{|\tilde{I}_G(h) - I_G(h)|}{|I_G(h)|} \rightarrow 0 \quad \text{for } h \rightarrow \infty.$$

This can be done in a similar way as in the proof of Lemma 1.23 in Marquardt [17]. In particular, since  $|I_G(h)| \geq |\tilde{I}_G(h)|$  it follows from (3.8) that

$$\frac{|\tilde{I}_1(h) - \tilde{I}_G(h)|}{|I_G(h)|} \leq \frac{|\tilde{I}_1(h) - \tilde{I}_G(h)|}{|\tilde{I}_G(h)|} \rightarrow 0.$$

For  $d < 0.5$  we get  $|I_G(h)| \geq |C_2|^2 \frac{h^{2d-1}}{1-2d}$  and for all  $h \geq K$ ,  $K$  large enough, we have  $|g_d(s + h)| \leq 2|C_2|h^{d-1}$ . There exists also a constant  $C_g > 0$  with  $\sup_{s \geq 0} |g_d(s)| \leq C_g$ . This yields for  $h \geq K$

$$\begin{aligned} |I_1(h) - \tilde{I}_1(h)| &= \left| \int_0^{h^{d/2}} g_d(s + h) g_d(s) \left[ M_2 + \sum_{m=2}^{\infty} \frac{(g_d(s))^{m-1}}{m!} M_{m+1} \right] ds \right| \\ &\leq \int_0^{h^{d/2}} 2|C_2|h^{d-1} C_g [M_2 + M^* C_g^{-1} (e^{C_g} - C_g - 1)] ds \\ &\leq 2|C_2| C_g [M_2 + M^* C_g^{-1} (e^{C_g} - C_g - 1)] h^{2d-1-d/2} \end{aligned}$$

and hence

$$\frac{|I_1(h) - \tilde{I}_1(h)|}{|I_G(h)|} \leq \frac{2C_g [M_2 + M^* C_g^{-1} (e^{C_g} - C_g - 1)] h^{2d-1-d/2}}{|C_2| \frac{h^{2d-1}}{1-2d}} \rightarrow 0 \quad \text{for } h \rightarrow \infty.$$

Similarly we get

$$\frac{|\tilde{I}_G(h) - I_G(h)|}{|I_G(h)|} \leq \frac{C_g h^{2d-1-d/2}}{|C_2| \frac{h^{2d-1}}{1-2d}} \quad \text{for } h \rightarrow \infty,$$

from which the result follows. Analogously we get with

$$\int_0^\infty g_d^k(s+h)g_d(s)ds \sim C_3 h^{(k+1)d-k}, \quad k \geq 2,$$

that

$$I_k(h) := \int_0^\infty \frac{1}{k!} g_d^k(s+h)g_d(s) \left[ M_{k+1} + \sum_{m=2}^\infty \frac{(g_d(s))^{m-1}}{m!} M_{m+k} \right] ds = o(h^{2d-1}).$$

and hence it follows that

$$\Psi_{d,\infty}^h(1) + \Psi_d^h(1) - 2\Psi_{d,\infty}(1) \sim C_1 h^{2d-1}, \quad \text{for } h \rightarrow \infty,$$

which proves the assertion.  $\square$

**Example 3.3.** In this example we consider a fractionally integrated ECOGARCH(1, 0.4, 1) process driven by a Lévy process  $L$  with Lévy symbol

$$\psi_L(u) = -\frac{u^2}{2} + \int_{\mathbb{R}} (e^{iux} - 1) \lambda \Phi_{0,1/\lambda}(dx),$$

where  $\Phi_{0,1/\lambda}(\cdot)$  is the distribution function of a normal distribution with mean 0 and variance  $1/\lambda$ . This means that  $L$  is the sum of a standard Brownian motion  $W$  and the compound Poisson process

$$J_t = \sum_{k=1}^{N_t} Z_k, \quad J_{-t} = \sum_{k=1}^{-N_{-t}} Z_{-k}, \quad t \geq 0,$$

where  $(N_t)_{t \in \mathbb{R}}$  is an independent Poisson process with intensity  $\lambda > 0$  and jump times  $(T_k)_{k \in \mathbb{Z}}$ . The Poisson process  $N$  is also independent from the i.i.d. sequence of jump sizes  $(Z_k)_{k \in \mathbb{Z}}$ , with  $Z_1 \sim N(0, 1/\lambda)$ . The Lévy process  $M$  is in this case given by the following expression

$$M_t = \sum_{k=1}^{N_t} [\theta Z_k + \gamma |Z_k|] - Ct, \quad t > 0,$$

with  $C = \gamma \int_{\mathbb{R}} |x| \lambda \Phi_{0,1/\lambda}(dx) = \sqrt{\frac{2\lambda}{\pi}} \gamma$ .  $M_{-t}$ ,  $t \geq 0$  is defined analogously. The parameter  $\theta$  is equal to  $-0.15$  and  $\gamma$  is equal to  $0.1$ . The stationary log-volatility process is of the form

$$\log(\sigma_{d,t+}^2) = \mu + \int_{-\infty}^t b_1 e^{-a_1(t-s)} dM_{d,s}, \quad t > 0,$$

where  $\mu = -5$ ,  $a_1 = 0.5$  and  $b_1 = 1$ . In Figure 1 we plotted the sample path of the simulated log-price process  $G_d$  observed at 30 000 equidistant time points, on the left hand side of the first row. On the right hand side one sees 6 000 observations of the return process  $G_d^{(1)}$ . The second row shows the empirical autocorrelation function of the volatility process  $\sigma_d^2$  and the corresponding sample path. In the last row the log-volatility process  $\log(\sigma_d^2)$  is described. The empirical autocorrelation function of the volatility process shows the same asymptotic behaviour as the empirical autocorrelation function of the log-volatility process.

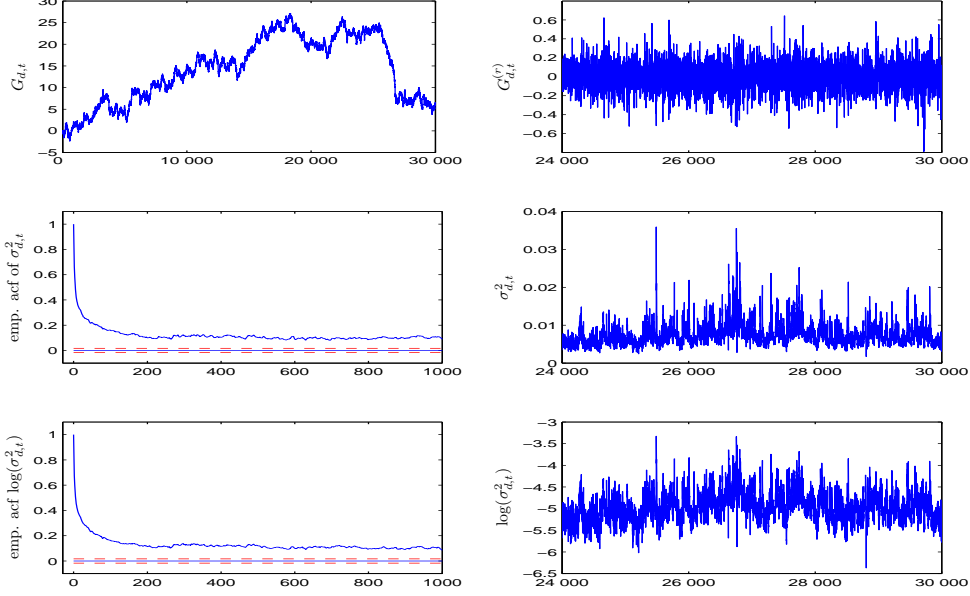


FIGURE 1: The log-price process  $G_d$  (top left) and 6000 observations of the return process  $G_d^{(1)}$  (top right) with parameters  $a_1 = 0.5, b_1 = 1, \mu = -5, \theta = -0.15, \gamma = 0.1$  and  $d = 0.4$ ; empirical autocorrelation function of  $\sigma_d^2$  (middle left) and 6000 observations of the volatility process  $\sigma_d^2$  (middle right); empirical autocorrelation function of  $\log(\sigma_d^2)$  (bottom left) and 6000 observations of the log-volatility process  $\log(\sigma_d^2)$  (bottom right). The jumps of the compound Poisson process are  $N(0, 1/2)$  distributed.

**Remark 3.4.** In the last Theorem we have shown, that the autocovariance function of the volatility process decays at a hyperbolic rate. For the discrete time EGARCH process this was shown by Surgailis and Viano [25].

In the continuous time setting Comte and Renault [11] showed this effect for a long memory stochastic volatility model, where the log-volatility process was modeled as an OU process driven by a fractional Brownian motion. Hence our result can also be applied to a continuous-time stochastic volatility model, where the log-price process  $Y = (Y_t)_{t \geq 0}$  satisfies

$$dY_t = \sigma_t dW_t, \quad t \geq 0, \quad (3.9)$$

with a Brownian motion  $W$ , and the log-volatility process  $\log(\sigma^2)$  is described by a FICARMA( $p, d, q$ ),  $p \geq q$ , process, where the Lévy measure of the driving noise process has finite moments of all orders  $k \geq 2$ .

#### 4. Second order properties of the return process

Second order properties are now derived under the assumption that the log-volatility process is strictly stationary. The structure of the price process  $G_d$  is the same as that of an ECOGARCH( $p, q$ ) process. Therefore the result concerning the first and second moment, as well as the autocovariance function, of the return process is analogously to the result in Proposition 5.1 in Haug and Czado [14].

**Proposition 4.1.** *Let  $L$  be a Lévy process with  $\mathbb{E}(L_1) = 0$  and  $\mathbb{E}(L_1^2) < \infty$ . Assume that  $\log(\sigma_d^2)$  is strictly stationary with marginal distribution  $F_d$ , where  $F_d$  is infinitely divisible with characteristic triplet  $(\gamma_{d,\infty}, 0, \nu_{d,\infty})$  and  $1 \in K_{d,\infty}$ . Then  $\mathbb{E}(G_{d,t}^2) < \infty$  for all  $t \geq 0$ , and for every  $t, h \geq r > 0$  it holds*

$$\mathbb{E}G_{d,t}^{(r)} = 0 \quad (4.1)$$

$$\mathbb{E}(G_{d,t}^{(r)})^2 = e^{\mu + \Psi_{d,\infty}(1)r} \mathbb{E}(L_1^2) \quad (4.2)$$

$$\text{Cov}(G_{d,t}^{(r)}, G_{d,t+h}^{(r)}) = 0. \quad (4.3)$$

If further  $\mathbb{E}(L_1^4) < \infty$  and the volatility process has finite second moment, then  $\mathbb{E}(G_{d,t}^4) < \infty$  for all  $t \geq 0$  and for every  $t, h \geq r > 0$  it holds

$$\text{Cov}((G_{d,t}^{(r)})^2, (G_{d,t+h}^{(r)})^2) = \mathbb{E}(L_1^2) \int_h^{h+r} \text{Cov}(G_{d,r}^2, \sigma_s^2) ds. \quad (4.4)$$

#### Acknowledgements

We like to thank Claudia Klüppelberg and Alexander Lindner for helpful comments and useful discussions concerning the COGARCH model. SH thanks Tina Marquardt for providing some insight into the long memory property of FICARMA processes. This work was supported by the Deutsche Forschungsgemeinschaft, Sonderforschungsbereich 386, Statistical Analysis of Discrete Structures.

#### References

- [1] Andersen TG and Bollerslev T. Heterogeneous Information Arrivals and Return Volatility Dynamics: Uncovering the Long-Run in High Frequency Returns. *Journal of Finance* 1997; **52**: 975-1005.
- [2] Anh VV, Heyde CC and Leonenko NN. Dynamic models of long-memory process driven by Lévy noise. *Journal of Applied Probability* 2002; **39**: 730-747.
- [3] Applebaum D. *Lévy Processes and Stochastic Calculus* 2004. Cambridge University Press.
- [4] Baillie RT. Long memory processes and fractional integration in econometrics. *Journal of Econometrics* 1996; **73**: 5-59.
- [5] Baillie RT, Bollerslev T and Mikkelsen HO. Fractionally integrated generalized autoregressive conditional heteroskedasticity. *Journal of Econometrics* 1996; **73**: 3-30.
- [6] Bollerslev T. Generalized Autoregressive Conditional Heteroskedasticity. *Journal of Econometrics* 1986; **31**: 307-327.
- [7] Bollerslev T and Mikkelsen HO. Modeling and pricing long memory in stock market volatility. *Journal of Econometrics* 1996; **73**: 151-184.



- [8] Brockwell PJ. Representations of continuous-time ARMA processes. *Journal of Applied Probability* 2004; **41A**: 375–382.
- [9] Brockwell PJ, Chadraa E and Lindner AM. Continuous time GARCH processes. *Ann. Appl. Probab.* 2006; **16**: 790–826.
- [10] Brockwell PJ and Marquardt T. Lévy-driven and fractionally integrated ARMA processes with continuous time parameter. *Statist. Sinica* 2005; **15**: 477–494.
- [11] Comte F and Renault E. Long memory in continuous-time stochastic volatility models. *Mathematical Finance* 1998; **8**: 291–323.
- [12] Engle RF. Autoregressive Conditional Heteroskedasticity With Estimates of the Variance of U.K. Inflation. *Econometrica* 1982; **50**: 987–1008.
- [13] Engle RF and Bollerslev T. Modeling the persistence of conditional variances. *Econometric Reviews* 1986; **5**: 1–50.
- [14] Haug S and Czado C. An exponential continuous time GARCH process. *Discussion paper 480, SFB386* 2006.
- [15] Kazakevičius V, Leipus R. A new theorem on existence of invariant distributions with applications to ARCH processes. *Journal of Applied Probability* 2003; **40**: 147–162.
- [16] Klüppelberg C, Lindner AM and Maller R. A continuous time GARCH process driven by a Lévy process: stationarity and second order behaviour. *Journal of Applied Probability* 2004; **41**: 601–622.
- [17] Marquardt T. Fractional Lévy processes, CARMA processes and related topics. PhD Thesis. Munich University of Technology.
- [18] Marquardt T. Fractional Lévy processes with an application to long memory moving average processes. *Bernoulli* 2006; to appear.
- [19] Mikosch T and Stărică C. Long range dependence effects and ARCH modelling. In *Theory and Applications of Long Range Dependence*, Doukhan P, Oppenheim G, and Taqqu Ms (eds) 2002: 439–459. Birkhäuser, Boston.
- [20] Nelson DB. ARCH models as diffusion approximations. *Journal of Econometrics* 1990; **45**: 7–38.
- [21] Nelson DB. Conditional heteroskedasticity in asset returns: a new approach. *Econometrica* 1991; **59**: 347–370.
- [22] Samko SG, Kilbas AA and Marichev OI. *Fractional Integrals and Derivatives* 1993. Gordon & Breach Science Publishers.
- [23] Sato KI. *Lévy Processes and Infinitely Divisible Distributions* 1999. Cambridge University Press.
- [24] Sato KI and Yamazato M. Stationary processes of Ornstein-Uhlenbeck type. *Probability Theory and Mathematical Statistics Lect. Notes in Math.* (ed. K. It and J. V. Prokhorov) ; **1021**: 541–551. Springer, Berlin.
- [25] Surgailis D and Viano MC. Long memory properties and covariance structure of the EGARCH model. *ESAIM Probability & Statistics* 2002; **6**: 311–329.